CS 747, Autumn 2023: Lecture 4

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Autumn 2023

Multi-armed Bandits

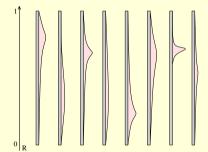
- The exploration-exploitation dilemma
- Definitions: Bandit, Algorithm
- e-greedy algorithms
- Evaluating algorithms: Regret
- Achieving sub-linear regret
- A lower bound on regret
- UCB, KL-UCB algorithms
- Thompson Sampling algorithm
- Understanding Thompson Sampling
- Concentration bounds
- Analysis of UCB
- Other bandit problems

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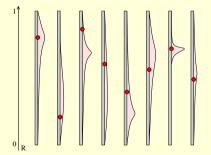
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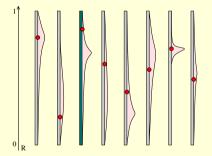


- Computational step: For every arm a, draw a sample

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• Bayes' Rule of Probability for events A and B:

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- Evidence samples e_1, e_2, \ldots, e_m are produced i.i.d. by the unknown world w.
- How to refine our belief distribution based on incoming evidence?

$$Belief_m(w) = \mathbb{P}\{w|e_1, e_2, \ldots, e_m\}.$$

 $Belief_{m+1}(w) = \mathbb{P}\{w|e_1, e_2, ..., e_{m+1}\}$

$$\begin{aligned} & \textit{Belief}_{m+1}(\textit{w}) = \mathbb{P}\{\textit{w}|\textit{e}_1,\textit{e}_2,\ldots,\textit{e}_{m+1}\} \\ &= \frac{\mathbb{P}\{\textit{e}_1,\textit{e}_2,\ldots,\textit{e}_{m+1}|\textit{w}\}\mathbb{P}\{\textit{w}\}}{\mathbb{P}\{\textit{e}_1,\textit{e}_2,\ldots,\textit{e}_{m+1}\}} \end{aligned}$$

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We achieve exactly that by taking

$$Belief_m(x) = Beta_{s+1,f+1}(x)dx$$

when the first *m* pulls yield *s* 1's and *f* 0's!

Principle of Selecting Arm to Pull

- We have a belief distribution for each arm's mean.
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- We sample a bandit instance I from the joint belief distribution, and
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- Alternative view: the probability with which we pick an arm is our belief that it is optimal. For example, if $A = \{1, 2\}$, the probability of pulling 1 is

$$\mathbb{P}\{x_1^t > x_2^t\} = \int_{x_1=0}^1 \int_{x_2=0}^{x_1} Beta_{s_1^t+1, f_1^t+1, (x_1)}Beta_{s_2^t+1, f_2^t+1, (x_2)}dx_2dx_1.$$

- 1. Understanding Thompson Sampling
- 2. Concentration bounds

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• Then, for or any fixed $\epsilon > 0$, we have

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 $\mathbb{P}\{\bar{x} \le \mu - \epsilon\} \le e^{-2u\epsilon^2}.$

• Note the bounds are trivial for large ϵ , since $\bar{x} \in [0, 1]$.

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We have *u* samples of *X*. How do we fill up this blank?:
 With probability at least 1 − δ, the empirical mean x̄ exceeds the true mean μ by at most ϵ₀ = _____.

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• We have u samples of X. How do we fill up this blank?: With probability at least $1 - \delta$, the empirical mean \bar{x} exceeds the true mean μ by at most $\epsilon_0 =$ _____.

We can write $\epsilon_0 = \sqrt{\frac{1}{2u} \ln(\frac{1}{\delta})}$; by Hoeffding's Inequality:

$$\mathbb{P}\{\bar{\mathbf{x}} \geq \mu + \epsilon_0\} \leq \mathbf{e}^{-2u(\epsilon_0)^2} \leq \delta.$$

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Yes. Assume u; $x_1, x_2, ..., x_u$; ϵ as defined earlier. Consider $Y = \frac{X-a}{b-a}$; for $1 \le i \le u$, $y_i = \frac{x_i-a}{b-a}$; $\bar{y} = \frac{1}{u} \sum_{i=1}^{u} y_i$. Since Y is bounded in [0, 1], we get

$$\mathbb{P}\{\bar{\mathbf{x}} \ge \mu + \epsilon\} = \mathbb{P}\left\{\bar{\mathbf{y}} \ge \frac{\mu - a}{b - a} + \frac{\epsilon}{b - a}\right\} \le e^{-\frac{2u\epsilon^2}{(b - a)^2}}, \text{ and}$$
$$\mathbb{P}\{\bar{\mathbf{x}} \le \mu - \epsilon\} = \mathbb{P}\left\{\bar{\mathbf{y}} \le \frac{\mu - a}{b - a} - \frac{\epsilon}{b - a}\right\} \le e^{-\frac{2u\epsilon^2}{(b - a)^2}}.$$

A "KL" Inequality

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• Then, for or any fixed $\epsilon \in [0, 1 - \mu]$, we have

$$\mathbb{P}\{ar{\mathbf{x}} \geq \mu + \epsilon\} \leq \mathbf{e}^{-\mathbf{u}\mathbf{KL}(\mu + \epsilon, \mu)},$$

and for or any fixed $\epsilon \in [0, \mu]$, we have

$$\mathbb{P}\{\bar{\mathbf{X}} \leq \mu - \epsilon\} \leq \mathbf{e}^{-u\mathsf{KL}(\mu - \epsilon, \mu)},$$

where for $p, q \in [0, 1]$, $KL(p, q) \stackrel{\text{\tiny def}}{=} p \ln(\frac{p}{q}) + (1 - p) \ln(\frac{1 - p}{1 - q})$.

Some Observations

The KL inequality gives a tighter upper bound:
 For *p*, *q* ∈ [0, 1],

$$\mathit{KL}(\rho,q) \geq 2(
ho-q)^2 \implies e^{-\mathit{u}\mathit{KL}(\rho,q)} \leq e^{-2\mathit{u}(
ho-q)^2}$$

- Both bounds are instances of "Chernoff bounds", of which there are many more forms.
- Similar bounds can also be given when *X* has infinite support (such as a Gaussian), but might need additional assumptions.

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