# Regret Minimisation in Multi-Armed Bandits Using Bounded Arm Memory

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#### **Abstract**

Regret minimisation in stochastic multi-armed bandits is a well-studied problem, for which several *optimal* algorithms have been proposed. Such algorithms depend on (sufficient statistics of) the empirical reward distributions of the arms to decide which arm to pull next. In this paper, we consider the design of algorithms that are constrained to store statistics from only a *bounded* number of arms. For bandits with a finite set of arms, we derive a sub-linear upper bound on the regret that decreases with the "arm memory" size M. For instances with a large, possibly infinite, set of arms, we show a sub-linear bound on the quantile regret.

Our problem formulation generalises that of Liau et al. (2018), who fix M=O(1), and so do not obtain bounds that depend on M. More importantly, our algorithms keep exploration and exploitation tightly coupled, without a dedicated exploration phase as employed by Liau et al. (2018). Although this choice makes our analysis harder, it leads to much-improved practical performance. For bandits with a large number of arms and no known structure on the rewards, our algorithms serve as a viable option. Unlike many other approaches to restrict the memory of bandit algorithms, our algorithms do not need any additional technical assumptions.

### 1 Introduction

We consider the well-studied problem of regret minimisation in stochastic multi-armed bandits (Robbins 1952; Berry and Fristedt 1985), which has applications in areas as diverse as drug-testing (Armitage 1960; Colton 1963), crowd-sourcing (Tran-Thanh et al. 2014), and on-line advertising (Li et al. 2010). Having to maximise its rewards by sequentially sampling unknown distributions, an agent must strike an optimal balance between exploring all the distributions and exploiting the most rewarding ones. Several algorithms have been proposed in the literature (Auer, Cesa-Bianchi, and Fischer 2002; Audibert and Bubeck 2009) with their loss (or regret) within a constant factor of optimal (Lai and Robbins 1985; Auer et al. 2003). Invariably, these algorithms maintain statistics of the pulls performed on each arm (such as the number of pulls and the observed mean reward); on every round, they determine the next arm to pull

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based on this empirical evidence. Hence, even if a constant number of statistics are recorded for each arm, the memory size needed to run such algorithms is  $\it linear$  in the number of arms. In this paper, we consider a setup in which algorithms may only store statistics from a fixed, specified, number of arms,  $M \geq 2$ .

Our setting is motivated by a number of practical applications, and is also of independent theoretical interest. In crowd-sourcing (Tran-Thanh et al. 2014) instead of tasking all the available workers simultaneously, an experimenter is often constrained to recruit a fixed number of workers for a given duration and assign jobs sequentially among them. In on-line advertising (Schwartz, Bradlow, and Fader 2017) an ad agency might have a large number of products to advertise, but the budget or space only to display a small, fixed number of ads for some specified period.

Researchers have long been drawn to the design of learning agents with a bounded memory size—starting with problems in hypothesis testing (Robbins 1956; Isbell 1959; Cover, Freedman, and Hellman 1976). In the context of bandits, Cover (1968) considers the realisation of an algorithm as a finite state machine, which for two-armed Bernoulli instances, asymptotically converges to the optimal average reward. The model we present in this paper does allow the memory size to depend on the horizon—but restricts that the memory be populated based on rewards from a bounded number of arms. To the best of our knowledge, the earliest analysis of such a constraint was undertaken by Herschkorn, Peköz, and Ross (1996). They present a non-recalling algorithm, which can store statistics from only a single arm (that is, M=1). For infinite bandit instances with Bernoulli arms, their algorithm maximises the almost sure average reward over an infinite horizon. Berry et al. (1997) improve these results when the sampling distribution over the infinite set of Bernoulli arms is uniform. Peköz (2003) demonstrates peculiarities that arise in this setting if the the reward distributions of the arms are not stochastically-ordered.

With M=1, an algorithm essentially has to begin afresh when it replaces the arm in its memory with a new one. Hence, the design of algorithms in this regime mainly focuses on an optimal schedule for sampling and erasing from memory. The algorithms are also restricted to 0-1 reward

distributions, which are intimately tied to decision making. Expanding the memory size to  $M \geq 2$  opens up a completely new dimension, since one or more "good" arms can be kept in memory as a safeguard against bad luck. Indeed Liau et al. (2018) have recently proposed UCBCONSTANTSPACE, an algorithm that achieves sub-linear regret for finite bandit instances when M=4. Unlike previous results, their bounds hold in finite time, and for arbitrary bounded reward distributions. Yet, their result is limited in three respects, all of which we address in this paper.

The algorithm proposed by Liau et al. (2018) follows an "explore-then-commit" paradigm, wherein a pure-exploration phase targeted at identifying a rewarding arm is followed by a phase to exploit the same arm. The primary technical contribution of Liau et al. (2018) is therefore a best arm identification algorithm that needs to remember the statistics of exactly four arms at any time. Whereas this approach greatly simplifies the analysis—and delivers sub-linear regret in finite bandits—it is known to be inferior to strategies that fully interleave exploration and exploitation (Agrawal and Goyal 2013; Auer, Cesa-Bianchi, and Fischer 2002). Our algorithm, UCB-M, adopts the latter approach, which, as expected, performs much better in practice. We also formally show a sub-linear regret bound for UCB-M.

Even if they restrict the arm memory size, Liau et al. (2018) require that their algorithm eventually sample *all* the arms in the bandit instance. In other words, they assume that there is sufficient *time* to pull all the arms even if there is not *space* to simultaneously monitor all the arms. In modern applications, the constraint is usually the opposite: one is apt to run out of time before running out of space. With an eye on practical relevance, we devise an algorithm, QUCB-M, which does not have to explore the entire set of arms. Naturally we cannot guarantee sub-linear regret for QUCB-M, as we do for UCB-M. However, we show a sub-linear bound on the quantile regret (Roy Chaudhuri and Kalyanakrishnan 2018)—which applies even when the set of arms is infinite.

We design our algorithms to work with an arbitrary arm memory size  $M \geq 2$ , given as an input parameter. Consequently we obtain upper bounds on the regret and quantile regret as functions of M. Such bounds are novel; the algorithm and analysis of Liau et al. (2018) assume M is fixed to 4. Hence, in a scenario where the experimenter allows only M < 4, UCBCONSTANTSPACE (Liau et al. 2018) will not be a choice. Alternatively, if the experimenter allows M > 4, UCBCONSTANTSPACE is again unable to take advantage of the increased arm memory.

In short, restricting arm memory adds an extra layer to the exploration-exploitation dilemma, which our algorithms negotiate effectively. To the best of our knowledge, the structure and analysis of UCB-M are novel, and at any rate distinct from UCBCONSTANTSPACE (Liau et al. 2018). In fact, UCB-M represents a whole spectrum of algorithms parameterised by the arm memory size. We formalise our problem and then present our algorithms.

#### **2 Problem Definition**

A bandit instance  $\mathcal{B} = (\mathcal{A}, \mathcal{D})$  comprises a set of arms  $\mathcal{A}$  and a set of reward distributions  $\mathcal{D}$ . Each arm  $a \in \mathcal{A}$ , when

pulled, generates a reward drawn i.i.d. from the corresponding reward distribution  $D_a \in \mathcal{D}$ , which we assume to be supported in [0,1]. The expected reward of arm  $a \in \mathcal{A}$  is given by  $\mu_a = \mathbb{E}_{r \sim D_a}[r]$ . The experimenter has no information regarding  $\mathcal{D}$ , and can only gather knowledge about it by pulling the arms. The experimenter's interaction with the bandit instance results in a history  $H_t$  at each round  $t \geq 1$ :  $H_t \stackrel{\mathrm{def}}{=} \{(a_i, r_i)\}_{i=1}^t$ , where,  $r_i \in [0, 1]$  is the reward produced at i-th step by pulling the arm  $a_i \in \mathcal{A}$ .

**Regret Minimisation.** Assuming  $\mu^* \stackrel{\text{def}}{=} \min\{y \in [0,1]: \forall a \in \mathcal{A}, \mu(a) \leq y\}$ , and a given horizon of pulls T, the (conventional) cumulative regret incurred by an algorithm is defined as

$$\mathbb{R}_T^* \stackrel{\text{def}}{=} T\mu^* - \sum_{t=1}^T \mathbb{E}[\mu_{a_t}],\tag{1}$$

where  $a_t$  is the arm pulled by the algorithm on round t. The expectation is taken over random rewards and possible randomisation introduced by the algorithm. In recent work, Roy Chaudhuri and Kalyanakrishnan (2018) introduce "quantile regret", an alternative measure of performance that is especially relevant to infinite-armed bandits. Under their definition, a problem instance  $\mathcal{I}=(\mathcal{B},P_{\mathcal{A}})$  consists of a bandit instance  $\mathcal{B}$  as well as a sampling distribution  $P_{\mathcal{A}}$  for selecting arms from  $\mathcal{A}$ . Letting  $\rho \in [0,1]$ , the  $(1-\rho)$ -th quantile of  $P_{\mathcal{A}}$  is given by  $\mu_{\rho}=\inf\{x\in[0,1]:\Pr_{a\sim P_{\mathcal{A}}}\{\mu_a\leq x\}\geq 1-\rho\}$ . Then, for a specified horizon T, the quantile regret for parameter  $\rho$ , or the " $\rho$ -regret", is given by

$$\mathbb{R}_T(\rho) \stackrel{\text{def}}{=} T\mu_\rho - \sum_{t=1}^T \mathbb{E}[\mu_{a_t}]. \tag{2}$$

If  $\mathcal{A}$  is infinite, one cannot hope to achieve sub-linear regret unless some side information is available regarding  $\mathcal{A}$  and/or  $\mathcal{D}$  (for example, that the mean reward is a linear function of features of the arms (Dani, Hayes, and Kakade 2008)). Yet, even without side information, one can achieve sub-linear  $\rho$ -regret for  $\rho>0$  (Roy Chaudhuri and Kalyanakrishnan 2018).

Word RAM Model. Constrained to use a bounded memory size, an algorithm can either curtail the precision with which individual arms' statistics are stored and/or store statistics only from a small number of arms. In this paper, we consider a model that allows individual arms' sufficient statistics to be stored exactly, but which limits the storage to a total of M such arms. We consider the "word RAM" model (Aho, Hopcroft, and Ullman 1974; Cormen et al. 2009), which considers a word as the unit of space. This model assumes that each of the input values and variables can be stored in O(1) word space. For finite bandit instances ( $|\mathcal{A}| < \infty$ ), assuming the horizon  $T \gg |\mathcal{A}|$ , we consider a word to consist of  $O(\log T)$  bits. Therefore, the O(M) words available to our algorithm result in a total space complexity of  $O(M \log(|\mathcal{A}|T))$  bits, where the additive  $\log |\mathcal{A}|$  bits are required to uniquely address the arms.

For infinite bandit instances  $(|\mathcal{A}| = \infty)$ , assuming that the experimenter needs to analyse  $\rho$ -regret for some  $\rho \in [0,1]$  and a horizon  $T \gg 1/\rho$ , we take the word size as  $O(\log(T/\rho))$  bits, which leads to a total space complexity of  $O(M\log(T/\rho))$  bits. In this case an algorithm encounters an arm only if it is sampled by  $P_{\mathcal{A}}$ . It does not need the memory to address every arm in  $\mathcal{A}$ . Roughly  $O(1/\rho)$  arms need to be sampled to get one from the top  $\rho$ -fraction of arms, and so we allow each sampled arm to be indexed using  $O(\log(1/\rho))$  bits. We refer to the set of arms stored in memory as the arm memory, whose cardinality is at most M. Before pulling a new arm—one that is not currently in the arm memory—the algorithm must first load it in memory, possibly by deleting an existing arm.

Our aim is to devise algorithms that minimise (conventional) regret and quantile regret while using an arm memory of size at most M; below we formalise these two problems.

**CR-M.** An algorithm  $\mathcal{L}$  is said to *solve* CR-M if for all bandit instances  $(\mathcal{A}, \mathcal{D})$  and  $M \geq 2$ : for a sufficiently large horizon T (not necessarily known beforehand),  $\mathcal{L}$  achieves  $\mathbb{R}_T^* \in o(T)$  using an *arm memory size* at most M.

**QR-M.** An algorithm  $\mathcal{L}$  is said to *solve* QR-M if for all problem instances  $(\mathcal{B}, P_{\mathcal{A}})$  and  $M \geq 2$ : for all  $\rho \in (0, 1]$  and sufficiently large horizon T (that depends on  $\rho$ , and may not necessarily known beforehand),  $\mathcal{L}$  achieves  $\mathbb{R}_T(\rho) \in o(T)$  using an *arm memory size* at most M.

In the following section, we outline the building blocks of our algorithms for solving CR-M and QR-M. These algorithms are then presented in detail in sections 4 and 5, respectively.

## 3 Key Intuitions

To get familiar with our approach, first consider our simpler problem: CR-M for finite bandit instances  $(|\mathcal{A}|<\infty).$  Assume  $M<|\mathcal{A}|;$  otherwise one may apply regular algorithms such as UCB1 (Auer, Cesa-Bianchi, and Fischer 2002). Also assume, to simplify our argument, that there is a unique optimal arm  $a^*\in\mathcal{A}.$ 

Minimising a Product of Probabilities. Intuitively, an algorithm that solves CR-M for finite bandit instances must ensure that the probability of pulling  $a^*$  is increased by progressively increasing at least one of two probabilities: (1) the probability that  $a^*$  is in arm memory, and (2) if  $a^*$  is in arm memory, the probability that it will be pulled more often than the other arms in arm memory. For any algorithm that achieves sub-linear regret, we can write, with T denoting horizon:  $\mathbb{R}_T^* \in o(T) \implies \lim_{T \to \infty} \mathbb{R}_T^*/T = 0 \implies \lim_{T \to \infty} \sum_{t=1}^T \Pr\{a_t = a^*\}/T = 1$ . Now, imposing the arm memory constraint, and letting  $X_t$  be the arm memory at t-th pull, we notice  $\{a_t = a^*\} \implies \{a^* \in X_t\}$ . Therefore,

$$\Pr\{a_t = a^*\} = \mathbb{E}_{X_t} \left[ \Pr\{a_t = a^* | a^* \in X_t\} \Pr\{a^* \in X_t\} \right].$$
 (3)

Given a bandit instance, the UCBCONSTANTSPACE algorithm of Liau et al. (2018) first solves a pure-exploration problem up to a horizon  $\bar{T}$  to maximise the quantity  $\Pr\{a^* \in X_t\}$  in the R.H.S. of Equation (3). Once the number of pulls

crosses  $\bar{T}$  (which is determined based on the accrued rewards), the algorithm chooses the arm with the highest empirical reward in  $X_t$  and assigns the remaining pulls to that arm. Therefore, for  $t > \bar{T}$ , it switches to a pure-exploitation mode, now looking to maximise  $\Pr\{a_t = a^* | a^* \in X_{\bar{T}}\}$ .

We adopt a more balanced exploration strategy, aiming to simultaneously increase  $\Pr\{a^* \in X_t\}$  and  $\Pr\{a_t =$  $a^*|a^* \in X_t$  throughout the sampling process—with no explicit switch at  $\bar{T}$ . Note that for the sake of sufficient exploration, an algorithm must not stick to the same arm memory for too long. However, while selecting new arms (not in the current arm memory), it must judiciously decide which arms in memory to replace. It is desirable that the empirically-best arms in memory do not get replaced—this would help retain  $a^*$  in memory if it is already present. Achieving this property amounts to minimising simple regret, which is defined for the t-th round to be  $\mu^* - \mathbb{E}[\mu_{a_t}]$  if  $a_t \in \mathcal{A}$  denotes the arm pulled by the algorithm (Bubeck, Munos, and Stoltz 2009). To design our algorithms, we take advantage of a result by Bubeck, Munos, and Stoltz (2009) on the relation between (conventional) regret and simple regret.

Minimising Simple Regret. To present our approach, first we recall the definition of a general forecaster (Bubeck, Munos, and Stoltz 2009), depicted in Figure 1. Given a set of arms  $\mathcal{A}$  as input, at each step t, possibly depending on history  $H_{t-1}$ , the forecaster selects an arm  $a_t$  by using a strategy called the "allocation strategy". By pulling  $a_t$  it receives a reward  $r_t$ . Then it executes a "recommendation strategy" that takes  $H_t$  as the input and outputs an arm  $b_t$ . The forecaster alternates between its allocation and recommendation strategies until some stopping condition is met.

```
t = 1, H<sub>0</sub> = {∅}.
While (stopping condition is not met){
1. Execute allocation strategy that possibly depending on H<sub>t-1</sub> selects an arm a<sub>t</sub>.
2. Pull the arm a<sub>t</sub>, and get a reward r<sub>t</sub>. Update H<sub>t</sub> = H<sub>t-1</sub> ∪ {(a<sub>t</sub>, r<sub>t</sub>)}.
3. Execute recommendation strategy that possibly depending on H<sub>t</sub> outputs an arm recommendation b<sub>t</sub>. Update t = t + 1.
}
```

Figure 1: A general forecaster, as defined by Bubeck, Munos, and Stoltz (2009).

In Figure 1, if  $b_t \equiv a_t$ , then the cumulative regret (Equation 1, defined using  $a_t$ ) of the forecaster is identical to the sum of the simple regret over time steps t (defined using  $b_t$ ). Theorem 1 by Bubeck, Munos, and Stoltz (2009) shows that the simple and the cumulative regret cannot be minimised all the way to optimality simultaneously. However, as a positive result, they present an upper bound on the simple regret for a number of forecasters (Bubeck, Munos, and Stoltz 2009, Table 1), one of which is defined below.

**UCB-MPA.** A forecaster, which on each round uses UCB1 (Auer, Cesa-Bianchi, and Fischer 2002) as the allocation strategy, and uses a recommendation strategy that outputs the <u>most played arm</u> (MPA) thus far, is denoted UCB-MPA.

On their attempt to establish a problem-dependent upper bound on the simple regret incurred by UCB-MPA, Bubeck, Munos, and Stoltz (2009) [Section 4.2, Theorem 3] obtain a problem-independent bound, which, interestingly, becomes valid for horizons that exceed a problem-dependent threshold. The problem-dependent bound they present closely resembles the problem-independent bound, and does not offer much additional insight. In this paper, we quote their *problem-independent* upper bound (Bubeck, Munos, and Stoltz 2009, Theorem 3) in Theorem 3.1 which becomes the cornerstone of our analysis of UCB-M.

**Theorem 3.1** (Theorem 3 in Bubeck, Munos, and Stoltz (2009)). Given a K-sized set of arms A as input, if UCB-MPA runs for a horizon of T pulls such that  $T \ge K(K+2)$ , then for some constant C > 0, its simple regret is at most  $C\sqrt{K\log T/T}$ .

In summary, although UCB1 was originally designed as a regret minimisation algorithm (Auer, Cesa-Bianchi, and Fischer 2002), it still performs reasonably well as an exploration strategy—in fact ideal for the exploration-exploitation balance we seek for CR-M. We choose UCB-MPA over other forecasters as it is easy to comprehend and leads to a simpler derivation. Next we present an algorithm for solving CR-M on finite instances.

## 4 Algorithm for Finite-armed Bandit Instances

We present UCB-M (which internally uses UCB-MPA) and establish a problem-independent upper-bound on its cumulative regret. We empirically compare UCB-M and its variations with the algorithm by Liau et al. (2018). In principle, one may replace the underlying call to UCB1 with any other allocation strategy, such as Thompson sampling (Agrawal and Goyal 2012) or Moss (Audibert and Bubeck 2009), as we do in our experiments.

## 4.1 Upper-Bounding Conventional Regret

Algorithm 1 describes UCB-M, which solves CR-M on finite bandit instances. We improve upon the contribution of Liau et al. (2018) in three aspects: (1) UCB-M is empirically seen to be more efficient even with M=2 (as opposed to M=4 for theirs) as it avoids pure-exploration-based elimination; (2) it scales with the arm memory size; (3) we present a problem-independent upper bound on its regret.

Given a finite set of arms  $\mathcal{A}(|\mathcal{A}|=K<\infty)$  and arm memory size M ( $2\leq M< K$ ), UCB-M approaches the task in phases. It breaks each phase into  $h_0=\lceil (K-1)/(M-1)\rceil$  sub-phases. Inside any phase w, at each sub-phase j, it runs UCB-MPA on an M-sized subset of arms  $S^{w,j}$  (the arm memory), assigns the recommended arm to  $\hat{a}$ , and forwards it to the next sub-phase. On the subsequent sub-phase (which might belong to the next phase), it chooses M-1 new arms from  $\mathcal{A}$ , along with the arm  $\hat{a}$  forwarded from the previous sub-phase, and repeat the previous steps. It is to be noted that the horizon spent on each sub-phase of a phase w is the same and is given by  $b_w$ . Also, for  $w \geq 2$ , the total horizon spent in phase w is given by  $h_0 b_w = 2h_0 b_{w-1}$ . To satisfy

the assumption in Theorem 3 of Bubeck, Munos, and Stoltz (2009), in the first phase, for each of the sub-phases, UCB-M chooses a horizon of  $b_1=M(M+2)$  pulls. For  $M\geq K$ , as the arm memory size is large enough, it effectively removes the memory constraint. Without the constraint, it is preferable to directly run UCB1 (Auer, Cesa-Bianchi, and Fischer 2002) on the whole instance, which will incur a lower regret. We handle this case within UCB-M. Theorem 4.1 upper-bounds the regret incurred by the algorithm.

```
Algorithm 1: UCB-M: For finite bandit instances.
```

```
Input: A: the set of K arms indexed by [K], M(\geq 2): Arm
          memory size.
1 if M > K then
    Run UCB1 on A until the horizon runs out.
3 else
       b_1 = M(M+2).// Initial horizon per
         sub-phase
5
        \hat{a}=1. // Initial arm recommendation
       w=1. \ // \ {\it Counts} \ {\it number} \ {\it of} \ {\it phases}
       h_0 = \lceil (K-1)/(M-1) \rceil.// The number of
         sub-phases in a phase
       Randomly shuffle the arm indices.
       while the horizon is not finished // start a new
         phase
        do
10
11
            for each sub-phase j = 1, \dots, h_0; if the horizon is
12
              not finished do
                S^{w,j} = {\min\{l+1,K\}, \cdots, \min\{l+1+1\}}
13
                 (M-1), K\}.
                l \stackrel{\text{def}}{=} \max_a \{S^{w,j} \setminus \hat{a}\} \text{ if }
14
                  \{\{\hat{a} \in S^{w,j}\} \land \{\hat{a} > 1 + \max_{a} \{S^{w,j} \setminus \hat{a}\}\}\}
                  else \max_a S^{w,j}.
                if \hat{a} \not\in S^{w,j} then
15
                    \tilde{S}^{w,j} = \{\hat{a}\} \cup S^{w,j} \setminus \{l\}.
16
                    l = l - 1.
17
                /* ALLOCATION STRATEGY */
                Run UCB1 on S^{w,j} for horizon of b_w pulls or
18
                  the remaining horizon; whichever is smaller.
                /* RECOMMENDATION STRATEGY */
                \hat{a} \stackrel{\text{def}}{=} \text{The most played arm in } S^{w,j}. // To be
19
                  forwarded to next sub-phase
            w=w+1. // Increment phase count
20
            b_w = 2 \cdot b_{w-1}. // Increment horizon per
21
              sub-phase
```

**Theorem 4.1.** Given as input a set of K arms A, with  $K \ge 3$ , and an arm memory size M, such that  $2 \le M < K$ , for a horizon of T pulls, with  $T > KM^2(M+2)$ , UCB-M incurs regret  $\mathbb{R}_T^* = O\left(KM + (K^{3/2}/M)\sqrt{T\log(T/KM)}\right)$ .

**Proof Summary.** We outline our proof strategy here and furnish the full proof in Appendix A (see supplemental material for all our appendices). For a given bandit instance with K arms, an arm memory size M, and horizon T, the total

number of phases is upper bounded by  $\lceil \log_2(2T/MK) \rceil$  (Lemma A.2). In each phase, UCB-M ensures that every arm gets stored in the arm memory at least once (Corollary A.3). This fact, along with Theorem 3.1 helps us upper-bound the difference between the means of any optimal arm  $a^* \in \mathcal{A}$ , and the best arm in the current arm memory (Lemma A.5). We note that the regret incurred in each sub-phase is linearly dependent on that difference (Equation (4)). As the budget spent on each sub-phase progressively doubles in each phase, that difference falls quickly towards zero, resulting a sub-linear growth in regret. On the other hand, in each sub-phase the incurred regret is sub-linear with respect to the optimal arm in the arm memory (Lemma A.7). Subsequently, we take a sum over all the sub-phases across all the phases to prove the theorem (lemmas A.6 and A.8).

## 4.2 Experimental Evaluation

We empirically compare UCB-M and its variations with the algorithm UCBCONSTANTSPACE (Liau et al. 2018). The use of UCB1 (Auer, Cesa-Bianchi, and Fischer 2002) as a subroutine in Algorithm 1 can be replaced with any other allocation strategy, which in effect will give rise to a different upper bound. We consider MOSS (Audibert and Bubeck 2009) and Thompson Sampling (Agrawal and Goyal 2013) in our experiments, and rename UCB-M to TS-M and MOSS-M, respectively. Everything else in Algorithm 1, including the recommendation strategy, is kept unchanged.

We run our set of experiments on three different kinds of K-armed Bernoulli instances earlier used by Jamieson and Nowak (2014):  $\mathcal{B}_L^K$ ,  $\mathcal{B}_{0.3}^K$ , and  $\mathcal{B}_{0.6}^K$ . The means of the arms in  $\mathcal{B}_L^K$  are equally spaced between  $\mu_1=0.99$ , and  $\mu_K=0.01$ . For  $\mathcal{B}_{0.3}^K$  and  $\mathcal{B}_{0.6}^K$  the mean of the i-th arm is given by  $\mu_i=0.01+\mu^*-(\mu^*-0.01)((i-1)/(K-1))^{\alpha}$  for  $\alpha=0.3$ , and 0.6, respectively.

For K = 100, Figure 2(a) compares the cumulative regret incurred by UCB-M, TS-M, and Moss-M for an arm memory size M=2, with the UCBCONSTANTSPACE algorithm of Liau et al. (2018). A comparison of cumulative regret and the number of pulls to individual arms (on instances with K=10) is presented in Figure 3. It is important to note that despite using a larger arm-memory size M=4, which is twice that of the others, UCBCONSTANTSPACE incurs a significantly higher regret. Due to its pure-exploration phase, UCBCONSTANTSPACE spends a prohibitively large number of pulls on sub-optimal arms, leading to a high regret. By contrast, our algorithms just make sure that at any instant, the expected difference between the mean of the optimal arm and the best arm in the current arm memory is not too large. Apparently, this difference increases with the subsequent subphases. However, UCB-M chooses the optimal arm in its arm memory at least once in any given phase, leading to a "reset" of the difference. On the other hand, this difference progressively reduces across phases due to doubling the horizon in each phase, explaining why UCB-M, TS-M, and Moss-M incur significantly lower regret.

As UCB-M can take advantage of a larger arm memory size, we compare the incurred regret as a function of M. Recall that under UCB1 (Auer, Cesa-Bianchi, and Fis-

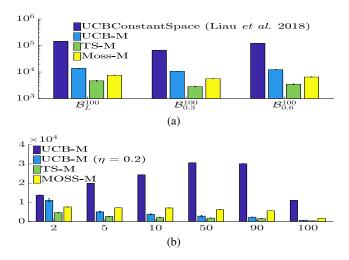


Figure 2: 2(a) presents a comparison of incurred cumulative regret in log scale (y axis), by UCB-M, TS-M, Moss-M, for M=2, and the UCBCONSTANTSPACE algorithm of Liau et al. (2018) (uses M=4) after  $10^6$  pulls. Each bar represents the average over 100 runs, with one standard error. For details about the instances, see Section 4.2. 2(b) presents a comparison of incurred regret (y axis) on the instance  $\mathcal{B}_L^{100}$  by different algorithms by varying arm memory size M (x axis), after  $10^6$  pulls. Each bar represents incurred regret averaged over 100 runs, with one standard error. For details about the instances, see Section 4.2.

cher 2002), if an arm a has been pulled  $u_a^t$  times up to round t, and if  $\hat{\mu}_a^t$  is its empirical average reward, then the upper confidence bound of that arm is given by  $ucb_a^t = \hat{\mu}_a^t + \eta \sqrt{2 \ln t/u_a^t}$ , with  $\eta = 1$ . It can be experimentally validated that tuning  $\eta$  can lead to achieving a smaller regret as claimed by the authors (Auer, Cesa-Bianchi, and Fischer 2002, Section 4). We present the regret incurred by UCB-M for  $\eta = 0.2$ , alongside the other algorithms.

Intuition suggests that increasing M should help in achieving a low regret, as it increases the chance of pulling the optimal arm more frequently. Also, the upper bound given by Theorem 4.1 supports this intuition. However, in practice, we notice an interesting behaviour. On the instance  $B_I^{100}$ , we compare the cumulative regret incurred by UCB-M, TS-M, and Moss-M by varying M. Figure 2(b) shows the results; for a comparison on other instances, the reader is referred to Figure 4 in Appendix B. As expected, UCB-M, TS-M and Moss-M always incur a higher regret than their unconstrained (M = K) counterparts. Also, for UCB-M with  $\eta = 0.2$ , TS-M and Moss-M, increasing the arm memory size M reduces the regret. However, the behaviour of UCB-M with  $\eta = 1$  is different. If M < K, it incurs a relatively low regret for M=2, afterwards increasing with M, followed by a slow decrease. On the other hand, UCB-M with  $\eta = 0.2$  and the other algorithms not only incur a lower regret but behave consistently. We posit that this peculiar behaviour is due to the looseness of the upper confidence bound in UCB1—which explains why it is absent if  $\eta = 0.2$ .

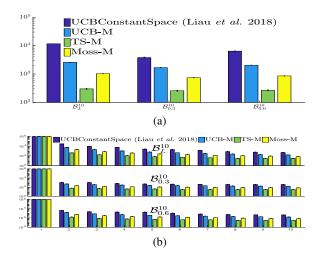


Figure 3: Comparison of incurred cumulative regret (3(a)) and the number of pulls to the individual arms (3(b)) in log scale (y axis), on the instances  $\mathcal{B}_L^{10}$ ,  $\mathcal{B}_{0.3}^{10}$  and  $\mathcal{B}_{0.6}^{10}$  by algorithms UCB-M, TS-M, Moss-M (for M=2), and UCBCONSTANTSPACE (Liau et al. 2018), after  $10^6$  pulls. Each bar represents the average over 100 iterations, and with one standard error. For details about the instances and the algorithms we refer to Section 4.2.

## 5 Algorithm for Infinitely-armed Bandit Instances

Quantile-regret minimisation is conceived by Roy Chaudhuri and Kalyanakrishnan (2018) as a practical approach in bandit instances in which the number of arms is too large for all to be sampled. It would be impossible to achieve sublinear regret on such instances unless some side information regarding the embedding of the arms or the rewards is available. By contrast, sublinear quantile regret can be obtained without such structural information: the simple idea is that if  $1/\rho$ arms are sampled at random, in expectation one will lie in the top  $\rho$  fraction. Roy Chaudhuri and Kalyanakrishnan (2018) have proven that for any  $\rho \in (0,1)$  (unknown to the algorithm), and for a sufficiently large horizon T, their algorithm QRM2 achieves  $\mathbb{R}_T(\rho) \in O(((1/\rho)\log(1/\rho))^{2.89} + T^{0.674})$ . In this section, we provide a bounded arm-memory algorithm, QUCB-M, that builds on this idea and enjoys sub-linear quantile regret. We empirically compare the algorithm with previous ones from the literature.

#### 5.1 Upper-Bounding Quantile Regret

We solve the QR-M problem by modifying the QRM2 algorithm (Roy Chaudhuri and Kalyanakrishnan 2018) to make it use UCB-M as the sub-routine, and adjust its arm exploration rate accordingly to minimise the upper bound. We call the resulting algorithm QUCB-M and describe it in Algorithm 2. Theorem 5.1 presents its upper bound on its quantile regret.

**Theorem 5.1.** Given as input any problem instance and an arm memory size  $M \geq 2$ : for  $\rho \in (0,1)$  and for a sufficiently large horizon T, the quantile regret incurred by QUCB-M is upper-bounded as  $R_T(\rho) \in$ 

#### Algorithm 2: QUCB-M

Input:  $\mathcal{A}, P_{\mathcal{A}}$ , and arm memory size M1 Set  $\alpha = 0.205, B = \left(M^2(M+2)\right)^{\frac{1}{1-\alpha}}$ , and  $\mathcal{K}_0 = \emptyset$ .

2 for  $r = 1, 2, 3, \cdots$  do

3  $t_r = 2^r B, \ n_r = \lceil t_r^{\alpha} \rceil$ .

4 Form a set  $\mathcal{K}_r$  by selecting additional  $n_r - |\mathcal{K}_{r-1}|$  arms from  $\mathcal{A}$  using  $P_{\mathcal{A}}$ , and adding to  $\mathcal{K}_{r-1}$ .

5 Run UCB-M on  $\mathcal{K}_r$ , for horizon of  $t_r$ , with arm memory size M.

$$o\left(\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{4.89} + MT^{0.205} + T^{0.81}\sqrt{\frac{\log M}{M^2}\log\frac{T}{M}}\right)$$

The proof applies Theorem 4.1 and follows the steps of the proof of Theorem 3.3 in the paper by Roy Chaudhuri and Kalyanakrishnan (2018). We present the detailed proof in Appendix C. Roy Chaudhuri and Kalyanakrishnan (2018) obtain the upper bound on the quantile regret by simultaneously minimising (1) the regret incurred before discovering a good arm, and once such an arm is discovered, (2) the regret incurred by the subsequent exploration-exploitation routine. The memory constraint on UCB-M means its regret upper bound (Theorem 4.1) grows faster than that of UCB1 as a function of the number of arms K. Since QUCB-M uses UCB-M as a subroutine, naturally, its optimised dependence on T comes out larger than that of QRM2.

It is to be noted that inside QUCB-M one can use the algorithm UCBCONSTANTSPACE by Liau et al. (2018) instead of UCB-M. However, as we have already shown in Section 4.2 that UCB-M is empirically superior to their algorithm, we do not consider this variation in our experiments.

#### **5.2** Experimental Evaluation

Although QUCB-M is designed with the aim of minimising quantile regret, we use conventional cumulative regret as the evaluation metric. Similar to UCB-M, the QUCB-M algorithm can be altered to use TS-M or Moss-M as the subroutine instead; we call the resulting variants QTS-M and QMoss-M, respectively. Algorithm 2 uses the value of  $\alpha$  that minimises the upper bound in Theorem 5.1. However, for empirical efficiency, we keep  $\alpha=0.347$  as done in QRM2 (Roy Chaudhuri and Kalyanakrishnan 2018).

We compare the regret of each of these algorithms with that of the algorithms by Herschkorn, Peköz, and Ross (1996) and Berry et al. (1997), and present the results in Table 1. We use the same four Bernoulli instances used by Roy Chaudhuri and Kalyanakrishnan (2018)—instances  $I_1$  and  $I_2$  have  $\mu^*=1$ , and the probability distributions on  $\mu$  induced by  $P_{\mathcal{A}}$  are given by  $\beta(0.5,2)$ , and  $\beta(1,1)$ , respectively. Similarly, instances  $I_3$  and  $I_4$  have  $\mu^*=0.6$ , and the probability distributions on  $\mu$  induced by  $P_{\mathcal{A}}$  are given by scaled  $\beta(0.5,2)$ , and  $\beta(1,1)$ , respectively. The rows of the table are labeled by the corresponding probability density function of encountering the mean rewards. As Table 1 shows, the existing algorithms incur a significantly higher regret in most of the cases. We put the comparison for  $\alpha=0.205$  in Table 2 in Appendix D.

Table 1: Comparison of cumulative regret  $(/10^5)$  of QUCB-M, QTS-M, QMoss-M (for  $\alpha=0.347$ ) with Non-Stationary algorithm (NS) proposed by Herschkorn, Peköz, and Ross (1996), and  $\sqrt{T}$ -run (STR),  $\sqrt{T} \ln T$ -learning (STLT), and Non-recalling  $\sqrt{T}$ -run (NST) algorithms of Berry et al. (1997) after  $10^6$  pulls, on instances  $I_1$ ,  $I_2$ ,  $I_3$  and  $I_4$ . Each result is the average of 20 runs, showing one standard error.

Algorithms	M	$I_1$ : $\beta(0.5, 2)$ $\mu^* = 1$	$I_2$ : $\beta(1,1)$ $\mu^* = 1$	$I_3$ : $\beta(0.5, 2)$ $\mu^* = 0.6$	$I_4$ : $\beta(1,1)$ $\mu^* = 0.6$
NS (1996)	1	3.58 ±0.4	1.11 ±0.2	$1.64 \pm 0.2$	$0.79 \pm 0.1$
STR (1997)	2	6.18±0.5	1.11±0.4	4.18±0.3	2.03±0.3
STLT (1997)	2	6.32±0.4	0.69±0.3	4.38±0.2	2.15±0.3
NST (1997)	1	5.35 ±0.5	<b>0.03</b> ±0.004	$4.56 \pm 0.001$	$2.55 \pm 0.001$
QUCB-M	2	1.84±0.17	0.41±0.02	1.29±0.10	0.49±0.02
	10	1.98±0.16	0.59±0.02	1.49±0.09	0.63±0.01
QUCB-M $(\eta = 0.2)$	2	2.00±0.20	0.32±0.05	1.41±0.10	0.69±0.04
	10	1.71±0.16	0.16±0.02	1.16±0.09	0.30±0.02
QTS-M	2	1.77±0.17	0.32±0.04	1.23±0.09	0.40±0.02
	10	1.91±0.16	0.18±0.03	1.14±0.10	0.30±0.02
QMoss-M	2	1.74±0.17	0.31±0.02	1.20±0.10	0.39±0.02
	10	<b>1.69</b> ±0.15	0.25±0.02	1.13±0.09	<b>0.30</b> ±0.010

It is interesting to note that like the finite instances, increasing arm memory leads to a lower regret. Specifically, the scaled version of QUCB-M (using UCB-M with  $\eta=0.2$ ), along with QTS-M and QMoss-M, show an improvement with larger arm memory. However, with  $\eta=1$  in the underlying UCB-M, QUCB-M fails to take the advantage of the larger arm memory.

### 6 Conclusion

We address the problem of regret minimisation in stochastic bandits using a bounded arm memory. This problem is relevant in applications where the number of arms is too large to store in memory, and there is no recourse to generalisation over the arms to reduce the number of parameters to learn. Most previous approaches have relied on additional assumptions such as a particular family of distributions for the mean reward, and the rewards being 0-1 (Herschkorn, Peköz, and Ross 1996; Berry et al. 1997). Liau et al. (2018) present an "explore-then-commit" algorithm that escapes such assumptions and still enjoys sub-linear regret. Yet, its "pure exploration" subroutine makes it inefficient in practice. The algorithm depends on eventually sampling every arm in the bandit instance, and so cannot be used on large or infinite instances—which are, in fact, the more relevant targets for the use of a bounded memory.

We propose UCB-M, an algorithm based on UCB1 (Auer, Cesa-Bianchi, and Fischer 2002) for finite bandit instances, which enjoys a sub-linear upper bound on regret, and is significantly more efficient in practice. Unlike previous algorithms, UCB-M offers the flexibility of using an arbitrary arm mem-

ory size, facilitating the experimenter to use as much memory as is available. In fact the unique design of UCB-M gives us the freedom of varying the arm memory size M: a clear contrast with UCBCONSTANTSPACE (Liau et al. 2018) which is designed specifically for M=4. Experiments validate our choice of using UCB1 in multiple phases instead of using an explore-then-commit approach: Figure 2(a) shows that UCB-M with a memory size of 2, exactly half that of UCBCONSTANTSPACE, already incurs less than a tenth of the regret on the instances and horizons tested. Additionally UCB-M can flexibly use other regret-minimisation routines other than UCB1 (Auer, Cesa-Bianchi, and Fischer 2002) (such as Thompson Sampling (Agrawal and Goyal 2012) and Moss (Audibert and Bubeck 2009)) while achieving a smaller regret, and retaining its theoretical guarantee. To the best of our knowledge the idea and the analysis of UCB-M are novel and do not resemble that of any existing algorithm.

We extend the algorithm UCB-M to QUCB-M, which uses QRM2 (Roy Chaudhuri and Kalyanakrishnan 2018) as a subroutine, and achieves sub-linear quantile regret under the bounded arm memory constraint. We empirically verify that QUCB-M incurs a lower conventional cumulative regret than existing constrained-memory algorithms (Herschkorn, Peköz, and Ross 1996; Berry et al. 1997) on a wide range of infinite bandit instances.

It would be interesting from a theoretical standpoint to take even the simpler among our problems—CR-M in *finite* bandits—and derive a non-trivial lower bound on the regret (note that  $\Omega(\sqrt{KT})$  carries over from the unconstrained setting (Auer et al. 2003)). The main technical challenge in so doing is to handle the arm memory—a random variable that limits algorithm's choices, while also being manipulated by the algorithm. To the best of our knowledge, related lower bounds have only been shown in special settings. For instance, Berry et al. (1997) show a lower bound on the failureproportion for infinitely-armed Bernoulli bandits under the uniform sampling distribution, while Peköz (2003) shows that non-recalling algorithms (M = 1) can have a linear lower bound unless the reward distributions of the arms are stochastically ordered. Liau et al. (2018) provide a conjecture on the lower bound for their problem, but a formal proof does not appear forthcoming. There is another line of work in the memory-restricted bandit setting by Lu and Lu (2011), who produce both upper and lower bounds on the regret. They restrict the number of rewards that can be remembered for each arm, rather than—as we restrict—the number of arms for which statistics can be maintained. Due to this fundamental difference, their techniques and bounds do not apply in our setting. Hence, establishing a lower bound for CR-M is an interesting challenge, which we leave for future work—while noting that such a bound would also be a prerequisite for a lower bound for QR-M.

## 7 Acknowledgement

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## A Proof of Theorem 4.1 from Section 4.1

We shall use  $\log$  and  $\ln$  to denote base 2 and natural logarithm respectively. Also, for any positive integer Z we shall denote the set  $\{1,2,\cdots,Z\}$  by [Z].

**Lemma A.1.** For a given K-sized set of arms A, and an arm memory size M < K, the number of sub-phases required to ensure that each arm in A has been chosen into arm memory at least once is not more than  $h_0$ .

*Proof.* We notice that at the beginning of each sub-phase there are exactly M-1 arms except the arm  $\hat{a}$  recommended from the previous step. Let, h be the maximum number of sub-phases possible in a phase. We realise that each phase w ends as soon as for every arm  $a \in \mathcal{A}$ , there exists a sub-phase j, such that  $S^{w,j} \ni a$ . Therefore,  $h = \min\{y : \mathcal{A} \subseteq \bigcup_{j=1}^y S^{w,j}\} = \left\lceil \frac{K-1}{M-1} \right\rceil = h_0$ .

**Lemma A.2.** For a given K-sized set of arms  $\mathcal{A}$ , and an arm memory size M < K, the number of phases UCB-M executes is upper bounded by  $x_0 \stackrel{\text{def}}{=} \left\lceil \log \frac{2T}{MK} \right\rceil$ .

*Proof.* Let x be the total number of phases executed by UCB-M. It should be noted that the value of M, K, and T might be such that the total horizon (T) runs out before finishing the last phase. Now, for any given phase w ( $w \geq 1$ ), the horizon spent on each sub-phase is the same, that is  $b_w = 2^{w-1}b_1$ . Therefore, we can write  $T = \sum_{w=1}^x \sum_{j=1}^{h_0} b_w = h_0b_1\sum_{w=1}^x 2^{w-1} = h_0b_1(2^x-1)$ . Hence, the number of phases (say x) can be upper bounded as

$$\begin{split} x &= \log \left( \frac{T}{h_0 b_1} + 1 \right) \leq \log \left( \frac{T}{\left\lceil \frac{K-1}{M-1} \right\rceil b_1} + 1 \right), \\ & \left[ \text{because}, h_0 = \left\lceil \frac{K-1}{M-1} \right\rceil \right], \\ & \leq \log \left( \frac{T}{\left\lceil \frac{K-1}{M-1} \right\rceil M(M+2)} + 1 \right) \\ & \left[ \text{because}, b_1 = M(M+2) \right], \\ & \leq \log \left( \frac{T}{\frac{K-1}{M-1} M(M+2)} + 1 \right) < \log \left( \frac{T}{KM} + 1 \right), \\ & \leq \log \left( \frac{2T}{KM} \right) \\ & \left[ \text{because}, T > 2MK \right], \\ & \leq \left\lceil \log \frac{2T}{MK} \right\rceil = x_0. \end{split}$$

For the rest of the analysis of UCB-M, we shall assume that the total number of phases executed by UCB-M is  $x_0$ . As UCB-M ensures inclusion of the optimal arm at least once in every phase, we note the following.

**Corollary A.3.** Let us denote, the sequence of sub-phasewise arm memory as  $\mathcal{S} \stackrel{\text{def}}{=} \{S^{1,1}, S^{1,2}, \cdots, S^{1,h_0}, S^{2,1}, S^{2,2}, \cdots, S^{2,h_0}, \cdots, S^{x_0,1}, S^{x_0,2}, \cdots, S^{x_0,h_0}\}$ . Then, for  $d \geq h_0$ , at least one of any d consecutive elements of  $\mathcal{S}$  contains  $a^*$ .

In any given phase, we need to upper-bound the difference of mean of the best in memory arm between two successive sub-phase. Considering  $\mathcal S$  as defined in Corollary A.3, let  $a_*^{y,j} \in S^{y,j} \in \mathcal S$  be the arm recommended by the sub-phase j-1 to j. It is important to note that  $\max_{a \in S^{y,j}} \mu_a \geq \mu_{a_*^{y,j}}$ . Therefore, in the interest of finding an upper bound on the regret, it is safe to consider  $\mu_{a_*^{y,j}} = \max_{a \in S^{y,j}} \mu_a$  as a pessimistic estimate of the best mean in  $S^{y,j}$ . In any given sub-phase  $j \leq h_0 - 1$  in phase y, we let  $\mathbb E[r^{y,j+1}] \stackrel{\text{def}}{=} \mathbb E[\mu_{a_*^{y,j}} - \mu_{a_*^{y,j+1}}]$ . Also, for any phase y, let  $\mathbb E[r^y]$  be a quantity such that, for all  $j \in \{1, \cdots, h_0 - 1\}$ ,  $\mathbb E[r^{y,j+1}] \leq \mathbb E[r^y]$ . Now, noticing that on each sub-phase in a phase y UCB-M spends exactly  $b_y$  pulls, we explicitly define  $\mathbb E[r^y]$  as follows.

**Corollary A.4.** In any phase y, at the end of each sub-phase j, the expected simple regret with respect to  $\mu_{a_*^{y,j}}$  is upperbounded by  $\mathbb{E}[r^y]$ . Hence, using Theorem 3.1, it is sufficient to define  $\mathbb{E}[r^y] \stackrel{\text{def}}{=} C\sqrt{(M \log b_y)/b_y}$ .

We notice that the arm forwarded from each sub-phase to the next one, not necessarily be the optimal arm. Hence, in the worst case, the expected difference between the mean of the optimal arm, and the highest mean reward in the current arm memory grows linearly with the number of sub-phase in a given phase. We upper bound it as follows.

**Lemma A.5.** Suppose, we are given a K-sized set of arms A, and an arm memory size M. Also, let at any phase  $y \geq 2$ , in the sub-phase j,  $\mu_*^{y,j}$  be the maximum of the mean of the arms in the arm-memory. Then  $\max_{1 \leq j \leq h_0} \mathbb{E}[\mu^* - \mu_*^{y,j}] \leq 2h_0 \mathbb{E}[r^y]$ .

*Proof.* Letting,  $\mathbb{E}[r_*^{y,j}] \stackrel{\text{def}}{=} \mathbb{E}[\mu^* - \mu_*^{y,j}]$ . We break the proof into two steps.

Step 1 upper bounds  $\mathbb{E}[r_*^{y,h_0}]$ , which is an upper bound on  $\mathbb{E}[r_*^{y,j}]$ , for all  $j \in [h_0]$ ; while Step 2 upper bounds  $\mathbb{E}[r_*^{y+1,j}]$ . Both the steps are based on Corollary A.3, that ensures at least one of the  $h_0$  consecutive sub-phases (not necessarily from the same phase) must contain the optimal arm  $a^*$  in the arm-memory.

**Step 1.** Let,  $1 \leq k_0 \leq h_0 - 1$  be the first sub-phase in phase y, to have  $a^*$  in the arm-memory. Therefore,  $k_0 \stackrel{\text{def}}{=} \min\{i \in [h_0]: a^* \in S^{y,i}\}$ , and hence, by definition,  $\mathbb{E}[r_*^{y,k_0+1}] = \mathbb{E}[r^{y,k_0+1}]$ . Therefore, for any subsequent sub-phase  $j \in \{k_0+1,\cdots,h_0\}$  in phase y,  $\mathbb{E}[r_*^{y,j}] = \mathbb{E}[\mu^* - \mu_*^{y,j}] = \mathbb{E}[\mu^* - \mu_*^{y,j}] = \mathbb{E}[\mu^* - \mu_*^{y,k_0+1}] + \sum_{v=k_0+2}^{j-1} \mathbb{E}[\mu_*^{y,v} - \mu_*^{y,v+1}]$ . As there are  $h_0$  sub-phases in any phase, hence, for all  $k_0+1 \leq j \leq h_0$ ,  $\mathbb{E}[r_*^{y,j}] \leq \mathbb{E}[r_*^{y,h_0}] \leq (h_0-k_0+1)\,\mathbb{E}[r^y] \leq h_0\,\mathbb{E}[r^y]$ .

**Step 2.** Let,  $j_0$  be the last sub-phase in phase y-1, such that  $a^* \in S^{y-1,j_0}$ . From Step 1,  $\mathbb{E}[r_*^{y-1,h_0}] \leq (h_0-j_0+1)\,\mathbb{E}[r^{y-1}]$ . Now, considering sub-phase i in phase y, we realise that if  $i\geq j_0$ , then there exists a sub-phase  $w\in\{1,\cdots,i\}$  such that  $a^*\in S^{y,w}$ . Now, for  $i\leq j_0-1$ ,

$$\begin{split} & E[r_*^{y,i}] \leq \max_{2 \leq j_0 \leq [h_0]} \max_{1 \leq i \leq j_0 - 1} \mathbb{E}[r_*^{y-1,h_0}] + i \cdot \mathbb{E}[r^y], \\ & \leq \max_{2 \leq j_0 [h_0]} \max_{1 \leq i \leq j_0 - 1} \leq (h_0 - j_0 + 1) \, \mathbb{E}[r^{y-1}] + i \cdot \mathbb{E}[r^y], \\ & \leq (h_0 - 1) \, \mathbb{E}[r^{y-1}] + \mathbb{E}[r^y] < 2h_0 \, \mathbb{E}[r^y], \\ & [\text{because } \mathbb{E}[r^y] \leq \mathbb{E}[r^{y-1}] \leq 2 \, \mathbb{E}[r^y]] \; . \end{split}$$

Together, Step 1 and Step 2 prove the lemma.  $\Box$ 

Next, we use Lemma A.5 to upper bound the cumulative regret  $\mathbb{R}_{T}^{*}$ .

**Bifurcation of**  $\mathbb{R}_T^*$ . For any given phase w, and a sub-phase j, let  $\mu_*^{w,j} \stackrel{\text{def}}{=} \max \mu_a : a \in S^{w,j}$ , and  $R_{w,j}$  be the incurred regret. Then,

$$R_{w,j}^* = b_w \mu^* - \sum_{t=1}^{b_w} \mathbb{E}[\mu_{a_t}]$$

$$= b_w \mathbb{E}[\mu^* - \mu_*^{w,j}] + \sum_{t=1}^{b_w} (\mathbb{E}[\mu_*^{w,j}] - \mathbb{E}[\mu_{a_t}]). \tag{4}$$

Where the expectation is taken over all possible sources of randomisation. Now, letting  $R_{w,j}^{(1)} = b_w(\mu^* - \mu_*^{w,j})$ , and  $R_{w,j}^{(2)} = \sum_{t=1}^{b_w} (\mathbb{E}[\mu_*^{w,j}] - \mathbb{E}[\mu_{a_t}])$ , we can write,

$$R_T^* = \sum_{w=1}^{x_0} \sum_{i=1}^{h_0} R_{w,j}^* = \sum_{w=1}^{x_0} \sum_{i=1}^{h_0} (R_{w,j}^{(1)} + R_{w,j}^{(2)}).$$
 (5)

Now, using Lemma A.5 we upper bound  $R_{w,j}^{(1)}$  as follows.

**Lemma A.6.** For  $2 \le M < K$ , and for  $T > KM^2(M + 2)$ , and for some constant C',  $\sum_{w=1}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(1)} \le C' \left(KM + (K^{3/2}/M)\sqrt{T\log(T/MK)}\right)$ .

Proof.

$$\begin{split} &\sum_{w=1}^{N_0} \sum_{j=1}^{h_0} R_{w,j}^{(1)} = \sum_{j=1}^{h_0} R_{1,j}^{(1)} + \sum_{w=2}^{N_0} \sum_{j=1}^{h_0} R_{w,j}^{(1)} \\ &\leq h_0 b_1 + \sum_{w=2}^{N_0} \sum_{j=1}^{h_0} R_{w,j}^{(1)} \leq h_0 b_1 + \sum_{w=2}^{N_0} \sum_{j=1}^{h_0} b_w \, \mathbb{E}[\mu^* - \mu_*^{w,j}] \\ &\leq h_0 b_1 + \sum_{w=2}^{N_0} \sum_{j=1}^{h_0} b_w (2h_0 \, \mathbb{E}[r^w]) \\ & [\text{using Lemma A.5}] \\ &\leq h_0 b_1 + 2C_1 h_0^2 \sum_{w=2}^{N_0} b_w \sqrt{\frac{M \log b_w}{b_w}} \, [\text{using Corollary A.4}] \\ &\leq h_0 b_1 + 2C_1 h_0^2 \sum_{w=2}^{N_0} \sqrt{b_w M \log b_w} \\ &\leq h_0 b_1 + 2C_2 h_0^2 \sqrt{M b_1} \sum_{w=2}^{N_0} \left( 2^{w-1} \log \left( 2^{w-1} b_1 \right) \right)^{\frac{1}{2}} \\ & [\text{because, } b_w = 2^{w-1} b_1] \\ &= h_0 b_1 + 2C_2 h_0^2 \sqrt{M b_1} \sum_{w=2}^{N_0} \left( (w-1 + \log b_1) \, 2^{w-1} \right)^{\frac{1}{2}} \\ &\leq h_0 b_1 + 2C_2 h_0^2 \sqrt{M b_1} \sum_{w=2}^{N_0} \left( (w-1) 2^{w-1} \right)^{\frac{1}{2}} \\ &\leq h_0 b_1 + C_3 h_0^2 \sqrt{M b_1} \sum_{w=2}^{N_0} \left( (w-1) 2^{w-1} \right)^{\frac{1}{2}} \\ &\leq h_0 b_1 + C_4 h_0^2 \sqrt{M b_1} \left( (w-1) 2^{w-1} \right)^{\frac{1}{2}} \\ &\leq C_5 \left( \left\lceil \frac{K-1}{M-1} \right\rceil M(M+2) + \\ \left\lceil \frac{K-1}{M-1} \right\rceil^2 \sqrt{M^2 (M+2)} \left( \frac{T}{MK} \log \frac{T}{MK} \right)^{\frac{1}{2}} \right) \\ &[\text{substituting for } b_1, h_0 \text{ and } x_0] \\ &\leq C_6 \left( \frac{K}{M} M^2 + \left( \frac{K}{M} \right)^2 \sqrt{M^3} \left( \frac{T}{MK} \log \frac{T}{MK} \right)^{\frac{1}{2}} \right) \\ &\leq C_7 \left( KM + \frac{K^{3/2}}{M} \sqrt{T \log \frac{T}{MK}} \right), \end{split}$$

wherein,  $C_1, C_2, \cdots, C_7$  are appropriate constants.  $\Box$ 

We note, that  $\sum_{w=1}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(2)}$ , can be upper-bounded using the problem independent upper-bound on the cumulative regret of UCB1 (Auer, Cesa-Bianchi, and Fischer 2002), as we restate below.

**Lemma A.7** (Problem-Independent Upper Bound on Cumulative Regret of UCB1 Auer, Cesa-Bianchi, and Fischer (2002)). Given a set of K-arms as the input, for any horizon T, the cumulative regret incurred by UCB1

 $R_T^* \le 12\sqrt{TK\log T} + 6K$ . Further, if  $T \ge K/2$ , then  $R_T^* \le 18\sqrt{TK\log T}$ .

Next, using Lemma A.7, we upper bound  $\sum_{w=1}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(2)}$ .

**Lemma A.8.** For  $2 \le M < K$ , and  $T > KM^2(M+2)$ , and for some constant C'' > 0,  $\sum_{w=1}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(2)} \le C'' \left(KM + \sqrt{TK \log(T/MK)}\right)$ .

*Proof.* We notice that at any sub-phase j of any phase  $w \ge 2$ , due to Lemma A.7, there exists a constant C, such that  $R_{w.j}^{(2)} \le C\sqrt{b_w M \log b_w}$ . Therefore,

$$\begin{split} \sum_{w=1}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(2)} &= \sum_{j=1}^{h_0} R_{1,j}^{(2)} + \sum_{w=2}^{x_0} \sum_{j=1}^{h_0} R_{w,j}^{(2)}, \\ &\leq h_0 b_1 + \sum_{w=1}^{x_0} \sum_{j=1}^{h_0} C \sqrt{b_w M \log b_w} \text{ [using Lemma A.7]}, \\ &= h_0 b_1 + h_0 C \sum_{w=1}^{x_0} \sqrt{2^{w-1} b_1 M \log (2^{w-1} b_1)}, \\ &= \left[ \sum_{w=1}^{K-1} M(M+2) + C_1 \left[ \frac{K-1}{M-1} \right] \sqrt{M^2 (M+2)} \cdot \right] \\ &\qquad \left( \sum_{w=1}^{x_0} 2^{w-1} \left( (w-1) + \log(M(M+2)) \right) \right) \\ &= \left[ \text{substituting for } h_0 \text{ and } b_1 \right], \\ &\leq C_2 \left( \frac{K-1}{M-1} M(M+2) + \frac{K-1}{M-1} \sqrt{M^2 (M+2)} \sum_{w=1}^{x_0} \sqrt{(w-1) 2^{w-1}} \right) \\ &\qquad \left[ \text{because, } T > K M^2 (M+2) \right] \\ &\leq C_3 \left( \frac{K}{M} M^2 + \frac{K}{M} M^{3/2} \sqrt{x_0 2^{x_0}} \right) \\ &= C_3 \left( K M + K \sqrt{M} \sqrt{\frac{T}{MK}} \log \frac{T}{MK} \right), \end{split}$$

wherein,  $C_1, C_2, C_3, C_4$  are appropriate constants.

**Proof of Theorem 4.1** We prove the theorem by plugging Lemma A.6 and Lemma A.8 in Equation 5.

## B Additional Experimental Results from Section 4.2

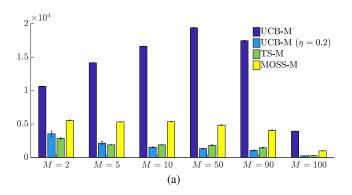


Figure 4: Comparison of incurred regret on the instance  $\mathcal{B}_{0.3}^{100}$ . Each bar represents regret incurred after  $10^6$  pulls, averaged over 100 iterations, and with one standard error. For details about the instances and the algorithms we refer to Section 4.2.

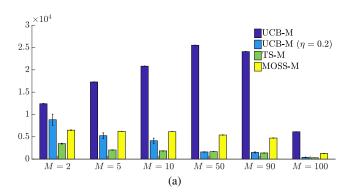


Figure 5: Comparison of incurred regret on the instance  $\mathcal{B}_{0.6}^{100}$ . Each bar represents regret incurred after  $10^6$  pulls, averaged over 100 iterations, and with one standard error. For details about the instances and the algorithms we refer to Section 4.2.

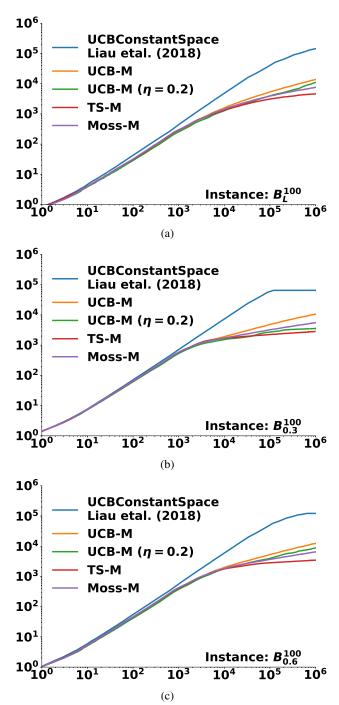


Figure 6: Comparison of incurred regret (along y-axis) on the instances  $\mathcal{B}_L^{100}$  (6(a)),  $\mathcal{B}_{0.3}^{100}$  (6(b)), and  $\mathcal{B}_{0.6}^{100}$  (6(c)), plotted over the number of pulls up to a horizon  $10^6$ , averaged over 100 runs. Both x and y axis are in  $\log_{10}$  scale. While the algorithm UCBConstantspace (Liau et al. 2018) uses M=4, all the other algorithms are run with M=2. For details about the instances and the algorithms we refer to Section 4.2.

### C Proofs from Section 5.1

In this appendix we provide the materials to complete the proof of Theorem 5.1. To prove the theorem we follow the steps of proof for Theorem 3.3 in Roy Chaudhuri and Kalyanakrishnan (2018). For any fixed  $\rho \in (0,1)$ , we break the analysis for upper bound on  $\mathbb{R}_T(\rho)$  in cases—first, the algorithm never encounters an arm from  $\mathcal{TOP}_{\rho}$ ; second, it picks at least one arm from  $\mathcal{TOP}_{\rho}$ . The key step in the analysis of the first part is showing that there exists  $r^* \geq 1$  such that for all  $r \geq r^*$ , the set of arms  $\mathcal{K}_r$  is sufficiently large to contain an arm from  $\mathcal{TOP}_{\rho}$  with high probability. Below, for any given  $\rho \in (0,1)$ , Lemma C.1 gives us an explicit form  $r^*$  and lower bounds  $\mathcal{K}_r$  in in terms of  $\rho$ .

**Lemma C.1.** Let,  $r^* = \left\lceil \frac{1}{\alpha} \log \left( \frac{1}{\rho} \log \frac{1}{\rho} \right) - \log B \right\rceil$ . Then, for every phase  $r \geq r^*$ , the size of  $\mathcal{K}_r$  can be lower bounded as  $n_r = \left\lceil t_r^{\alpha} \right\rceil \geq \left\lceil \frac{\alpha \log e}{(1+\gamma)\rho} \cdot \ln t_r \right\rceil$ , wherein,  $0.53 < \gamma \stackrel{\text{def}}{=} \max_x \frac{\log \log x}{\log x} < 0.531$ .

*Proof.* We notice, for every,  $r \geq r^*$ ,  $t_r \geq \left\lceil \left(\frac{1}{\rho} \log \frac{1}{\rho}\right)^{\frac{1}{\alpha}} \right\rceil$ . Then, for each  $r \geq r^*$ , we can lower bound the size of the set  $\mathcal{K}_r$  as follows. As,  $|\mathcal{K}_r| = n_r = \left\lceil t_r^{\alpha} \right\rceil$  is an integer, to ease the calculation let us define  $s_u = 2^u B$ , where  $u \in \mathbb{R}^+$ , and therefore,  $s_u \in \mathbb{R}^+$  does not need to be an integer. Now,

letting  $u^* \stackrel{\text{def}}{=} \log \left( \frac{1}{\rho} \log \frac{1}{\rho} \right)^{\frac{1}{\alpha}}$ , we get

$$\frac{1}{\rho} \log s_{u^*}$$

$$= \frac{1}{\rho} \log \left(\frac{1}{\rho} \log \frac{1}{\rho}\right)^{\frac{1}{\alpha}},$$

$$= \frac{1}{\alpha \rho} \left(\log \frac{1}{\rho} + \log \log \frac{1}{\rho}\right),$$

$$= \frac{1}{\alpha \rho} \log \frac{1}{\rho} \left(1 + \frac{\log \log \frac{1}{\rho}}{\log \frac{1}{\rho}}\right),$$

$$\leq \frac{1+\gamma}{\alpha \rho} \log \frac{1}{\rho}, \left[\operatorname{as} \gamma \stackrel{\text{def}}{=} \max_{x} \frac{\log \log x}{\log x}\right]$$

$$= \frac{1+\gamma}{\alpha} \left(\frac{1}{\rho} \log \frac{1}{\rho}\right),$$

$$= \frac{1+\gamma}{\alpha} s_{u^*}^{\alpha}.$$

$$\Rightarrow s_{u^*}^{\alpha} \geq \frac{\alpha \log e}{(1+\gamma)\rho} \log s_{u^*}.$$
(6)

As,  $s_u^{\alpha}$  grows with u faster than  $\log s_u$ , therefore,

$$\forall u \ge u^*, \ s_{u^*}^{\alpha} \ge \frac{\alpha \log e}{(1+\gamma)\rho} \ln s_{u^*}$$

Therefore, recalling that r is an integer, for all values of  $r \geq \lceil r^* \rceil$ , the statement of the lemma follows.  $\square$ 

Now, to show that for any phase  $r \geq r^*$  the probability of not having at least one arm from the set  $\mathcal{TOP}_{\rho}$  is sufficiently low. To this end, we define an event that no arm from  $\mathcal{TOP}_{\rho}$  is in the set  $\mathcal{K}_r$  as  $E_r(\rho) \stackrel{\text{def}}{=} \{\mathcal{K}_r \cap \mathcal{TOP}_{\rho} = \emptyset\}$  and upper bound its probability in the following Lemma C.2.

**Lemma C.2.** Assuming  $r^*$  and  $\gamma$  as defined in Lemma C.1, for any given  $\rho \in (0,1)$ , and for any  $r \geq r^*$ ,  $\Pr\{E_r(\rho)\} \leq t_r^{-\alpha \log e/(1+\gamma)}$ .

*Proof.* We note  $\Pr\{E_r(\rho)\}=(1-\rho)^{n_r}$ . Now, for some  $\alpha\in(0,1)$  that shall be tuned later, let  $r^*=\lceil(1/\alpha)\log((1/\rho)\log(1/\rho))\rceil$ . Therefore, in the round  $r^*$ , the number of pulls is given by  $t_{r^*}=2^{r^*}=((1/\rho)\log(1/\rho))^{1/\alpha}$ . Now, for  $r\geq r^*$ , the number of arms in  $\mathcal{K}_r$  is given by  $n_r=t_r^\alpha\geq\lceil(\alpha/((1+\gamma)\rho))\cdot\ln t_r^{\log e}\rceil$ , wherein,  $\gamma=\max_x(\log\log x)/\log x$   $(0.53<\gamma<0.531)$ .

Therefore, 
$$\Pr\{E_r(\rho)\} = (1 - \rho)^{n_r} \le \exp(-\lceil (\alpha/((1 + \gamma))) \cdot \ln t_r^{\log e} \rceil) \le t_r^{-\alpha \log e/(1+\gamma)}$$
.

Lemma C.2 shows that beyond the phase  $r^*$  the set  $\mathcal{K}_r$  will contain an arm from  $\mathcal{TOP}_\rho$  with high probability. Now, assuming there exists an arm in  $\mathcal{K}_r$  from  $\mathcal{TOP}_\rho$ , for each  $r \geq r^*$ , we can upper bound the accumulated regret beyond the  $r^*$ -th phase as follows.

**Lemma C.3.** Let, for  $r^*$  defined in Lemma C.1, given that for all  $r \geq r^*$ , algorithm QUCB-M has encountered at least one arm from  $T\mathcal{OP}_{\rho}$ , that is  $\mathcal{K}_r \cup T\mathcal{OP}_{\rho} \neq \emptyset$ . Then, the incurred regret beyond the round  $r^*$  is not more than  $C'\left(MT^{\alpha} + \frac{\sqrt{\log M}}{M}\sqrt{T^{1+3\alpha}\log\frac{T}{M}}\right)$ ; for some constant C'.

Proof.

$$\begin{split} &\sum_{r=r^*}^{\log(T/B)} C\left(n_r M + \frac{1}{M}\sqrt{n_r^3 t_r \log\frac{t_r}{n_r M}}\right), \\ &\sum_{r=r^*}^{\log(T/B)} C\left(t_r^\alpha M + \frac{1}{M}\sqrt{t_r^{1+3\alpha}\log\frac{t_r}{n_r M}}\right), \\ &= \sum_{\log(T/B)}^{\log(T/B)} C\left(2^{\alpha r} B^\alpha M + \frac{1}{M}\sqrt{2^{(1+3\alpha)r}B^{(1+3\alpha)}\log\left(\frac{B}{M}2^{(1-\alpha)r}\right)}\right), \\ &= \sum_{r=r^*}^{\log(T/B)} C\left(2^{\alpha r} B^\alpha M + \frac{\sqrt{B^{(1+3\alpha)}\log(B/M)}}{M}\sqrt{(1-\alpha)r}2^{(1+3\alpha)r}\right), \\ &\leq \sum_{r=r^*}^{\log(T/B)} C_1\left(2^{\alpha r} B^\alpha M + \frac{\sqrt{B^{(1+3\alpha)}\log(B/M)}}{M}\sqrt{r}2^{(1+3\alpha)r}\right) \\ &\leq \sum_{r=r^*}^{\log(T/B)} C_1\left(\left(\frac{T}{2^j B}\right)^\alpha B^\alpha M + \frac{\sqrt{\log(T/B)}}{M}\sqrt{\left(\frac{T}{2^j}\right)^{1+3\alpha}\log\frac{T}{B}}\right), \\ &= \sum_{j=0}^{\log(T/B)-r^*} C_2\left(M\left(\frac{T}{2^j}\right)^\alpha + \frac{\sqrt{\log M}}{M}\sqrt{\left(\frac{T}{2^j}\right)^{1+3\alpha}\log\frac{T}{B}}\right), \\ &\leq C_3\left(MT^\alpha\sum_{j=0}^{\log(T/B)-r^*}\left(\frac{1}{2^\alpha}\right)^j + \frac{\sqrt{\log M}}{M}T^{(1+3\alpha)/2}\sqrt{\log\frac{T}{B}}\sum_{j=0}^{\log(T/B)-r^*}\sqrt{\left(\frac{1}{2^j}\right)^{1+3\alpha}}\right), \\ &\leq C_4\left(MT^\alpha + \frac{\sqrt{\log M}}{M}\sqrt{T^{1+3\alpha}\log\frac{T}{B}}\right), \\ &\leq C_5\left(MT^\alpha + \frac{\sqrt{\log M}}{M}\sqrt{T^{1+3\alpha}\log\frac{T}{M}}\right), \\ &\leq C_7\left(MT^\alpha + \frac{\sqrt{\log M}}{M}\sqrt{T^{1+3\alpha}\log\frac{T}{M}}\right). \\ &\leq C_7\left(MT^\alpha + \frac{\sqrt{\log M}}{M}\sqrt{T^{1+3\alpha}\log\frac{T}{M}}\right).$$

for some constants  $C_1, C_2, C_3, C_4$  and  $C_5$ .

We note that Lemma C.3 is valid only if  $\mathcal{TOP}_{\rho} \cup \mathcal{K}_r \neq \emptyset$ 

for every  $r \geq r^*$ . However, there is a positive probability of never encountering any arm  $\mathcal{TOP}_{\rho}$ , to put formally,  $\mathcal{TOP}_{\rho} \cup \mathcal{K}_r = \emptyset$ , for all  $r \geq 1$ . Below, in Lemma C.4 we upper bound the accumulated regret in this case.

**Lemma C.4.** The expected regret due to not encountering any arm from the set  $\mathcal{TOP}_{\rho}$  is during running of the algorithm, is in  $O\left(\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}}+T^{1-\frac{\alpha\log e}{1+\gamma}}\right)$ .

*Proof.* Using Lemma C.1, we obtain the following.

$$\begin{split} & \sum_{r=1}^{\log(T/B)} t_r \Pr\{E_r(\rho)\} \\ & = \sum_{r=1}^{r^*-1} t_r \Pr\{E_r(\rho)\} + \sum_{r=r^*}^{\log(T/B)} t_r \Pr\{E_r(\rho)\} \\ & \leq \sum_{r=1}^{r^*-1} t_r + \sum_{r=r^*}^{\log(T/B)} t_r^{1 - \frac{\alpha \log e}{1 + \gamma}} \\ & \leq t_r^* + \sum_{r=r^*}^{\log(T/B)} t_r^{1 - \frac{\alpha \log e}{1 + \gamma}} \\ & = t_{r^*} + \sum_{r=r^*}^{\log(T/B)} (B2^r)^{1 - \frac{\alpha \log e}{1 + \gamma}} \\ & \leq t_{r^*} + B^{1 - \frac{\alpha \log e}{1 + \gamma}} \sum_{j=1}^{\log(T/B) - r^*} \left(\frac{T}{B2^j}\right)^{1 - \frac{\alpha \log e}{1 + \gamma}} \\ & \leq B \cdot 2^{\left\lceil \log\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}} - \log B\right\rceil} + \\ & T^{1 - \frac{\alpha \log e}{1 + \gamma}} \sum_{j=0}^{\log T - r^*} \left(\frac{1}{2}\right)^{j(1 - \frac{\alpha \log e}{1 + \gamma})} \\ & < 2^{\log\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}} + 1} + T^{1 - \frac{\alpha \log e}{1 + \gamma}} \sum_{j=0}^{\infty} \left(\frac{1}{2}\right)^{j(1 - \frac{\alpha \log e}{1 + \gamma})} \\ & = O\left(2^{\log\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}}} + T^{1 - \frac{\alpha \log e}{1 + \gamma}}\right) \\ & = O\left(\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}} + T^{1 - \frac{\alpha \log e}{1 + \gamma}}\right) \\ & = O\left(\left(\frac{1}{\rho}\log\frac{1}{\rho}\right)^{\frac{1}{\alpha}} + T^{1 - \frac{\alpha \log e}{1 + \gamma}}\right). \end{split}$$

**Proof of Theorem 5.1.** Together Lemma C.3 and Lemma C.3 establish the upper bound on  $\mathbb{R}_T(\rho)$ . Now, to minimise the upper bound we prove the theorem by noticing the upper bound is minimised with respect to T we set  $\alpha = 1/(3+2\log e/(1+\gamma)) \approx 0.205$ , and prove the theorem.

# D Additional Experimental Results from Section 5.2

Table 2: Cumulative regret  $(/10^5)$  of QUCB-M, QTS-M, QMoss-M and the strategies proposed by (Herschkorn, Peköz, and Ross 1996) and (Berry et al. 1997) after  $10^6$  pulls, on instances  $I_1$ ,  $I_2$ ,  $I_3$  and  $I_4$ . Each result is the average of 20 runs, showing one standard error.

Algorithms		$I_1: \beta(0.5, 2)$ $\mu^* = 1$	$I_2: \beta(1, 1)$ $\mu^* = 1$	$I_3$ : $\beta(0.5, 2)$ $\mu^* = 0.6$	$I_4: \beta(1, 1)$ $\mu^* = 0.6$				
Non-stationary Policy (Herschkorn, Peköz, and Ross 1996)		3.58 ±0.4	1.11 ±0.2	$1.64\pm0.2$	$0.79 \pm 0.1$				
$\sqrt{T}$ -run (Berry et al. 1997)	2	6.18±0.5	1.11±0.4	4.18±0.3	2.03±0.3				
$\sqrt{T} \ln T$ -learning (Berry et al. 1997)	2	6.32±0.4	0.69±0.3	4.38±0.2	2.15±0.3				
Non-recalling $\sqrt{T}$ -run (Berry et al. 1997)	1	5.35 ±0.5	0.03 ±0.004	$4.56\pm0.001$	$2.55 \pm 0.001$				
QUCB-M	2	3.69±0.34	0.74±0.11	2.27±0.21	0.51±0.07				
	10	4.26±0.37	0.91±0.19	2.65±0.22	0.63±0.11				
QUCB-M $\eta = 0.2$		3.67±0.35	0.72±0.12	2.21±0.21	0.55±0.08				
,	10	4.15±0.36	0.79±0.19	2.51±0.22	0.54±0.11				
QTS-M		3.14±0.39	0.62±0.07	1.97±0.19	0.44±0.07				
		3.88±0.35	0.67±0.13	2.49±0.23	0.45±0.06				
QMoss-M		3.64±0.34	0.70±0.11	2.21±0.21	0.46±0.07				
		4.16±0.36	0.80±0.19	2.53±0.22	0.52±0.11				

For  $\alpha=0.205$  the algorithms explore very small number of arms, that causes incorporating a good arm very unlikely leading to a high regret.