Quiz 1

15 Marks, 45 minutes

Thursday 25th August, 2016

Please answer to the point in the limited space provided for each question. You can do rough work in a separate sheet of paper provided to you. You can also assume any result stated or proved in the class (but NOT as part of the tutorials).

Problem 1. Let $\mathcal{D} = \langle (\mathbf{x}_1, y_1), \dots, (\mathbf{x}_m, y_m) \rangle$ such that each $y_j \in \Re$. Let $\phi(\mathbf{x}) = [\phi_1(\mathbf{x}), \dots, \phi_n(\mathbf{x})]$ be a vector of basis functions. Consider the linear regression function $f(\mathbf{x}) = \phi^T(\mathbf{x})\mathbf{w}$ with \mathbf{w} obtained either as a least squares or ridge regression estimate. Show that, using either of these estimates for \mathbf{w} , the regression function can be written in the (so-called *kernelized*)

form $f(\mathbf{x}) = \sum_{i=1}^{m} \alpha_i K(\mathbf{x}, \mathbf{x}_i) y_i$ where $K(\mathbf{x}, \mathbf{x}_i) = \phi^T(\mathbf{x}) \phi(\mathbf{x}_i)$ is a function of \mathbf{x} and \mathbf{x}_i only and independent of any of the \mathbf{y}_i 's and \mathbf{x}_j for all $j \neq i$. Each α_i can be a function of the entire dataset \mathcal{D} .

Hint: Use the following Matrix Identity that holds for any matrices P, B and R with compatible dimensions such that R and $BPB^T + R$ are invertible:

$$(P^{-1} + B^T R^{-1} B)^{-1} B^T R^{-1} = P B^T (B P B^T + R)^{-1}$$

(8 Marks)

Answer: The solution to linear (set $\lambda = 0$) and ridge regression can be written as $\mathbf{w} = (\Phi^T \Phi + \lambda I)^{-1} \Phi^T \mathbf{y}$ where

• Recall for Ridge Regression: $\mathbf{w} = (\Phi^T \Phi + \lambda I)^{-1} \Phi^T y$, where,

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \dots \\ y_m \end{bmatrix}$$

$$\Phi = \begin{bmatrix} \phi_1(\mathbf{x}_1) & \dots & \phi_p(\mathbf{x}_1) \\ \dots & \dots & \dots \\ \phi_1(\mathbf{x}_m) & \dots & \phi_p(\mathbf{x}_m) \end{bmatrix}$$

• Please note the difference between Φ and $\phi(\mathbf{x})$

$$\phi(\mathbf{x}_j) = \begin{bmatrix} \phi_1(\mathbf{x}_j) \\ \dots \\ \phi_p(\mathbf{x}_j) \end{bmatrix}$$

Then, the regression function will be

$$f(\mathbf{x}) = \phi^T(\mathbf{x})\mathbf{w} = \phi^T(\mathbf{x})(\Phi^T \Phi + \lambda I)^{-1}\Phi^T \mathbf{y}$$

• $\phi^T(\mathbf{x}_i)\phi(\mathbf{x}_j) = K(\mathbf{x}_i, \mathbf{x}_j)$

•
$$\left(\Phi^T \Phi\right)_{ij} = \sum_{k=1}^m \phi_i(\mathbf{x}_k) \phi_j(\mathbf{x}_k)$$

• $\left(\Phi\Phi^{T}\right)_{ij} = \sum_{k=1}^{p} \phi_k(\mathbf{x}_i)\phi_k(\mathbf{x}_j) = \phi^{T}(x_i)\phi(x_j) = K(\mathbf{x}_i, \mathbf{x}_j)$

Kernelizing Ridge Regression

• Given $\mathbf{w} = (\Phi^T \Phi + \lambda I)^{-1} \Phi^T \mathbf{y}$ and using the identity $(P^{-1} + B^T R^{-1} B)^{-1} B^T R^{-1} = PB^T (BPB^T + R)^{-1}$

$$- \Rightarrow$$
 by setting $R = I$, $P = \frac{1}{\lambda}I$ and $B = \Phi$,

- $\Rightarrow \mathbf{w} = \Phi^T (\Phi \Phi^T + \lambda I)^{-1} \mathbf{y} = \sum_{i=1}^m \alpha_i \phi(\mathbf{x}_i)$ where $\alpha_i = \left((\Phi \Phi^T + \lambda I)^{-1} \mathbf{y} \right)_i$
- \Rightarrow the final decision function $f(\mathbf{x}) = \phi^T(\mathbf{x})\mathbf{w} = \sum_{i=1}^m \alpha_i \phi^T(\mathbf{x})\phi(\mathbf{x}_i)$

The Kernel function in Ridge Regression

- We call $\phi^{\top}(x_1)\phi(x_2)$ a kernel function: $K(x_1, x_2) = \phi^{\top}(x_1)\phi(x_2)$
- The preceding expression for decision function becomes $f(\mathbf{x}) = \sum_{i=1}^{m} \alpha_i K(\mathbf{x}, \mathbf{x}_i)$ where $\alpha_i = (([K(\mathbf{x}_i, \mathbf{x}_j)] + \lambda I)^{-1} \mathbf{y})_i$

Problem 2. Case for non-IID dataset:

In the class, we discussed the case of Bayesian estimation for a univariate Gaussian from dataset \mathcal{D} that consisted of IID (independent and identically distributed) observations.

Let $\Pr(X) \sim \mathcal{N}(\mu, \sigma^2)$ and let σ^2 be known. Suppose, the examples $x_1...x_m$ in the dataset \mathcal{D} were not necessarily independent and whose possible dependence was expressed by known covariance matrix Ω but with a common unknown (to be estimated) mean $\mu \in \Re$. Let $\mathbf{u} = [1, 1, ..., 1]$ a *m*-dimensional vector of 1's and $\mathbf{x} = [x_1...x_m]$ and

$$Pr(x_1...x_m;\mu,\Omega) = \frac{1}{(2\pi)^{\frac{m}{2}}|\Omega|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x}-\mu\mathbf{u})^T\Omega^{-1}(\mathbf{x}-\mu\mathbf{u})}$$

Assume that $\Omega \in \Re^{m \times m}$ is positive-definite. Now answer the following questions

- 1. How would you go about doing Bayesian estimation for μ ? What will be an appropriate conjugate prior? What will the posterior be? And what will be the MAP and Bayes estimates?
- 2. Is the case of IID data set \mathcal{D} a special case of this problem? Prove your claim.

(7 Marks)

Answer:

This problem is directly adapted from Tutorial 2.

Answers to 1: As hinted in the class, we will expect the conjugate prior of mean μ of the (product of) Gaussian to be Gaussian. Let $\mu \sim \mathcal{N}(\mu_0, \sigma_0^2)$ with a fixed and known σ_0^2 .

$$\mathcal{N}(\mu_m, \sigma_m^2) = exp\left(\frac{-1}{2\sigma_m^2}(\mu - \mu_m)^2\right) = \Pr(\mu|\mathcal{D}) \propto \Pr(\mathcal{D}|\mu) \Pr(\mu) = \frac{1}{(2\pi)^{\frac{m}{2}}|\Omega|^{\frac{1}{2}}} \frac{1}{\sqrt{2\pi\sigma_0^2}} exp\left(-\frac{1}{2}(\mathbf{x} - \mu\mathbf{u})^T \Omega^{-1}(\mathbf{x} - \mu\mathbf{u}) - \frac{(\mu - \mu_0)^2}{2\sigma_0^2}\right) \propto exp\left(-\frac{1}{2}(\mathbf{x} - \mu\mathbf{u})^T \Omega^{-1}(\mathbf{x} - \mu\mathbf{u}) - \frac{(\mu - \mu_0)^2}{2\sigma_0^2}\right)$$

Our reference equality is:

$$exp\left(-\frac{1}{2}(\mathbf{x}^{T}\Omega^{-1}\mathbf{x}-2\mu\mathbf{x}^{T}\Omega^{-1}\mathbf{u}+\mu^{2}\mathbf{u}^{T}\Omega^{-1}\mathbf{u})-\frac{(\mu-\mu_{0})^{2}}{2\sigma_{0}^{2}}\right)=exp\left(\frac{-1}{2\sigma_{m}^{2}}(\mu-\mu_{m})^{2}\right)$$

Matching coefficients of μ^2 , we get

$$\frac{-\mu^2}{2\sigma_m^2} = \frac{-1}{2}\mu^2 \mathbf{u}^T \Omega^{-1} \mathbf{u} + \frac{-\mu^2}{2\sigma_0^2} \Rightarrow \frac{1}{\sigma_m^2} = \frac{1}{\sigma_0^2} + \mathbf{u}^T \Omega^{-1} \mathbf{u}$$

Matching coefficients of μ , we get

$$\frac{2\mu\mu_m}{2\sigma_m^2} = \mu \left(\mathbf{x}^T \Omega^{-1} \mathbf{u} + \frac{2\mu_0}{2\sigma_0^2} \right) \Rightarrow \mu_m = \sigma_m^2 \left(\mathbf{x}^T \Omega^{-1} \mathbf{u} + \frac{\mu_0}{\sigma_0^2} \right) \Rightarrow \frac{1}{1 + \sigma_0^2 \mathbf{u}^T \Omega^{-1} \mathbf{u}} \left(\sigma_0^2 \mathbf{x}^T \Omega^{-1} \mathbf{u} + \mu_0 \right)$$

 μ_m will be the MAP estimate of μ .

One can easily verify that setting $\Omega = I$ gives the IID case. In fact, one can also verify that the with $\Omega = I$, one gets the same Bayesian estimates as in the IID case discussed in the class.