

Lecture 3: Regression

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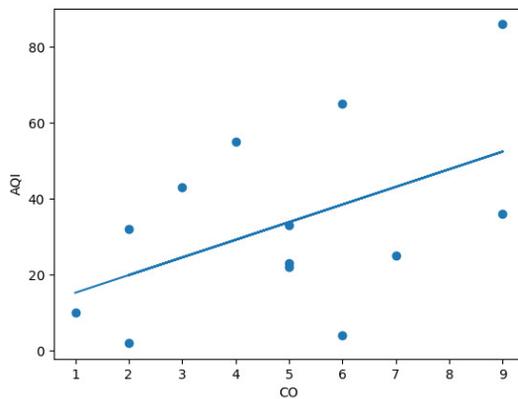
3.1 Introduction

Regression analysis finds a practical application in everyday life through the *Air Quality Index (AQI)*, a tool designed to quantify air pollution. Given the challenges of directly measuring and calculating the percentages of each air component like SO_2 and CO , regression comes into play

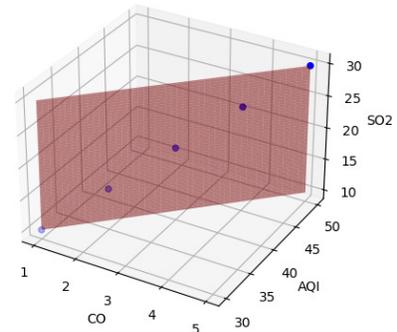
Instead of individually measuring all components, a subset is measured, and the remaining components are estimated using regression analysis. This statistical technique enables the establishment of relationships between measured and unmeasured components, offering a more efficient means of interpreting air quality by inferring the percentages of various pollutants without the need for exhaustive measurements.

$$AQI = \max\{f_1(x_1), f_2(x_2), \dots, f_n(x_n)\}$$

Where $f_i(x_i)$ is a unique function for each pollutant



(a) AQI vs CO



(b) AQI vs SO2 and CO

Figure 3.1: function of various pollutant concentration

In this case device might find only 1 or 2 functions and based on that guess the AQI i.e. to fit the perfect AQI data with limited observations and estimate AQI value.

3.2 Linear Regression

We use *Linear Regression* for estimating an unknown data from a know data as

1. It is a simple and powerful tool
2. It is interpretable
3. It's works on transformations of raw data

How do we best fit the given data?

We need a measurement criteria to calculate goodness of our estimation function. So we use an *error function* also called as *loss function*, *lost function*, *energy function*. It has two parameters *estimation function* and *data points*

$$\text{Error function} = E(f, D)$$

f is the estimation function

$$D = \{(x_1, y_1), (x_2, y_2), \dots (x_n, y_n)\} \text{ where } (x_i, y_i) \text{ is a data point from the data}$$

3.2.1 Possible Error functions

$$\sum_{i=1}^n (f(x_i) - y_i) \tag{3.1}$$

- 3.1 is not a good error function as it is signed.

$$\sum_{i=1}^n |f(x_i) - y_i| \tag{3.2}$$

- 3.2 is a better error function than 3.1 as it is unsigned.

$$\sum_{i=1}^n (f(x_i) - y_i)^2 \tag{3.3}$$

- 3.3 is squared cost function, most used error function

$$\sum_{i=1}^n (f(x_i) - y_i)^3 \tag{3.4}$$

- 3.4 is not a good error function as it is signed

3.2.2 Squared loss function

$$\sum_{i=1}^n (f(x_i) - y_i)^2$$

1. It is a *continuous function* and in particular *differentiable*.
2. Easy to visualize in Euclidean space.
3. Mathematical analysis become easier.

Let DS be the data set

$$DS = \{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)\}$$

Where x_i is the input and y_i is the output for the i^{th} training example. The number n = number of data samples or more usually called training instances. $x_i \in \mathbb{R}^d$ and $y_i \in \mathbb{R}$. Here \mathbb{R}^d is the d -dimensional space.

Let PM 2.5, SO_2 , CO be the components of x_i then,

$$x_i = \begin{bmatrix} x_{i_1} \\ x_{i_2} \\ x_{i_3} \end{bmatrix}$$

x_{i_1} represents the concentration of PM 2.5

x_{i_2} represents the concentration of CO

x_{i_3} represents the concentration of SO_2

Let us define a X matrix containing x_i

$$X = \begin{bmatrix} x_{11} & x_{12} & x_{13} & \dots & x_{1d} \\ x_{21} & x_{22} & x_{23} & \dots & x_{2d} \\ x_{31} & x_{32} & x_{33} & \dots & x_{3d} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ x_{n1} & x_{n2} & x_{n3} & \dots & x_{nd} \end{bmatrix} = \begin{bmatrix} x_1^T \\ x_2^T \\ x_3^T \\ \vdots \\ x_n^T \end{bmatrix}_{n \times d}$$

$$y = \begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{bmatrix}_{n \times 1}$$

3.2.3 General Regression

The goal of this is to find a function $f^*(x)$ such that it is the first prediction of y (output data) w.r.t. D .

$$f^* \in \arg \min E(f, D)$$

subject to $f \in \mathcal{F}$, where \mathcal{F} is the set of all functions

3.2.4 Parameterized Regression

In this f is a function of the form $f(x, w)$, where w are the parameters of regression.

e.g. $f(x, (\alpha, \lambda)) = \alpha e^{\lambda^T x}$

e.g. $f(x, w) = \sum w_i x_i$ i.e., $f(x, w) = w_0 + w_1 x + w_2 x^2 + \dots + w_k x^k$

We can use parameterized regression to reduce the solution space we needed to search from in case of a general regression. Let us take an example of parameterized function:

- $f(x, (\alpha, \lambda)) = \alpha e^{-\lambda^T x}$

In parameterized regression we need to minimize the parameterized function w.r.t the given parameters i.e.,

$$f \equiv f(x, w)$$

$$\arg \min_w E(f(x, w), D)$$

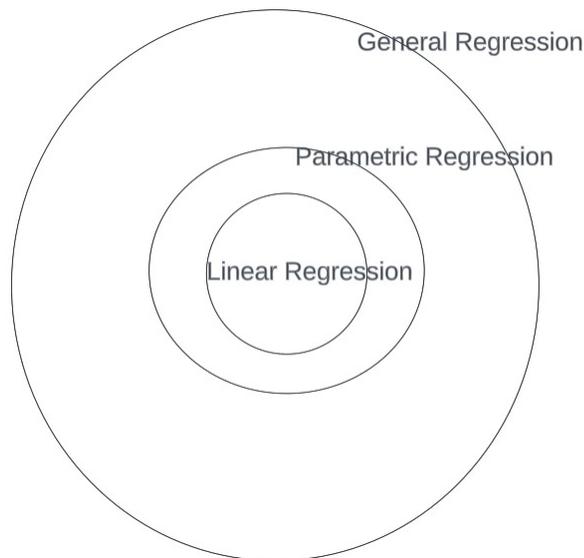


Figure 3.2: Diagram Representation

3.2.5 Linear Regression

In this f is a function of the form $f(x, w) = w^T x + w_0 = \bar{w}^T x$ here $w \in \mathbb{R}^d$

3.3 Least Square Optimisation for Linear Regression

$$W^* \in \arg \min_w \left(\sum_{i=1}^n \left(\sum_{j=0}^d w_j x_{ij} - y_i \right)^2 \right)$$

For $d = 1$

$$E(w, D) = \sum_{i=1}^n (y_i - w_0 - w_1 x_i)^2$$

$$\frac{\partial E}{\partial w_0} \implies -2 \left(\sum_{i=1}^n (y_i - w_0 - w_1 x_i) = 0 \right)$$

$$\implies \sum y_i - n w_0 - w_1 \sum x_i = 0$$

$$\frac{\partial E}{\partial w_1} \implies -2 \left(\sum_{i=1}^n x_i (y_i - w_0 - w_1 x_i) = 0 \right)$$

$$\implies \sum x_i y_i - w_0 \sum x_i - w_1 \sum x_i^2 = 0$$

Note that the 2 equations above are a linear equations in variables w_0 and w_1 . Solving for these we get,

$$w_1 = \frac{n * \sum x_i y_i - \sum x_i \sum y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

$$w_0 = \frac{\sum y_i \sum x_i^2 - \sum x_i \sum x_i y_i}{n \sum x_i^2 - (\sum x_i)^2}$$

3.3.1 Case 1: d=1

$$E(w, d) = \sum_{i=1}^n (y_i - w_0 - w_1 x_i)^2$$

Find w_0 w_1 such that

$$\frac{\partial E}{\partial w_0} = 0 \tag{3.5}$$

$$\frac{\partial E}{\partial w_1} = 0 \tag{3.6}$$

From equation 3.5

$$w_0 = \frac{\sum y_i - w_1 \sum x_i}{n} \tag{3.7}$$

From equation 3.6

$$w_1 = \frac{\sum x_i y_i - w_0 \sum x_i}{\sum x_i^2} \tag{3.8}$$

Take $\alpha = \frac{\sum x_i y_i}{\sum x_i^2}$ and $\beta = \frac{\sum x_i^2}{n}$

let \bar{x} be $\frac{\sum x_i}{n}$ and \bar{y} be $\frac{\sum y_i}{n}$

$$w_1 = \alpha - w_0 * \frac{\bar{x} * n}{n * \beta} \quad (3.9)$$

$$w_1 = \alpha - (\bar{y} - w_1 \bar{x}) \frac{\bar{x} * n}{n \beta} \quad \text{From 3.7} \quad (3.10)$$

$$w_1 * (1 - \frac{\bar{x}^2}{\beta}) = \alpha - \frac{\bar{y} \bar{x}}{\beta} \quad (3.11)$$

$$w_1 = \frac{\alpha * \beta - \bar{x} * \bar{y}}{\beta - \bar{x}^2} \quad (3.12)$$

Exercise: Find $\frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$ in terms of α and β (It turns out to be w_1)

3.3.2 Case 2: For d-dimensional data

$$x_i = \begin{bmatrix} x_{i1} \\ x_{i2} \\ \vdots \\ x_{id} \end{bmatrix} \quad X = \begin{bmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_n^T \end{bmatrix} \quad y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

Let $z_i = y_i - w^T x_i$

$$w^* \in \arg \min_w \sum_{i=1}^n (y_i - w^T x_i)^2 = \sum_{i=1}^n z_i^2 = \|z\|^2$$

$$z = \begin{bmatrix} z_1 \\ z_2 \\ \vdots \\ z_n \end{bmatrix} = \begin{bmatrix} y_1 - x_1^T w \\ y_2 - x_2^T w \\ \vdots \\ y_n - x_n^T w \end{bmatrix} = y - Xw$$

$$\|z\|^2 = \|y - Xw\|^2$$

$$w^* \in \arg \min_w \|y - Xw\|^2$$

$$\|y - Xw\|^2 = (Xw - y)^T (Xw - y)$$

$$\arg \min_w (Xw - y)^T (Xw - y) = (w^T X^T - y^T) (Xw - y)$$

$$E(w, D) = w^T X^T Xw - w^T X^T y - y^T Xw + y^T y$$

$$(w^T X^T y = y^T Xw)$$

$$E(w, D) = w^T X^T X w - 2y^T X w + y^T y$$

We can find w by doing $\nabla_w E = 0$

$$\text{So } \frac{\partial(2y^T X w)}{\partial w} = (2y^T X)^T = 2X^T y$$

$$\frac{\partial(w^T X^T X w)}{\partial w} = X^T X w + (X^T X)^T w = X^T X w + X^T X w = 2X^T X w$$

$$\frac{\partial(y^T y)}{w} = 0$$

$$\nabla_w E = 0 \implies 2X^T X w - 2X^T y = 0$$

$$\implies 2X^T X w = 2X^T y$$

$$\implies \boxed{w = (X^T X)^{-1} X^T y}$$

$$\boxed{(w^*)^T x = \hat{y}}$$

If $X^T X$ is not invertible then it means that the closed form expression cannot be used to find the optimal w^* . A possible such scenario is when there are less data points than the dimension of the data.